

Predictive Analytics

Ch5. The forecasters' toolbox

Prof. Dr. Benjamin Buchwitz



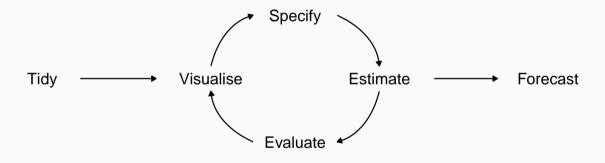
Outline

1 A tidy forecasting workflow

- 2 Some simple forecasting methods
- 3 Residual diagnostics
- 4 Distributional forecasts and prediction intervals
- 5 Forecasting with transformations
- 6 Forecasting and decomposition
- 7 Evaluating forecast accuracy

The process of producing forecasts can be split up into a few fundamental steps.

- 1 Preparing data
- 2 Data visualisation
- 3 Specifying a model
- 4 Model estimation
- 5 Accuracy & performance evaluation
- 6 Producing forecasts



Data preparation (tidy)

```
gdppc <- global_economy %>%
 mutate(GDP_per_capita = GDP/Population) %>%
  select(Year, Country, GDP, Population, GDP_per_capita)
gdppc
```

```
## # A tsibble: 15,150 x 5 [1Y]
## # Key: Country [263]
##
      Year Country
                              GDP Population GDP_per_capita
     <dbl> <fct>
                            <dbl>
                                      <dbl>
                                                     <dbl>
##
##
  1 1960 Afghanistan 537777811.
                                    8996351
##
   2 1961 Afghanistan 548888896.
                                    9166764
      1962 Afghanistan 546666678.
##
   3
                                    9345868
##
      1963 Afghanistan 751111191.
                                    9533954
                                                     78.8
   4
##
   5 1964 Afghanistan 800000044.
                                    9731361
                                                     82.2
      1965 Afghanistan 1006666638.
##
   6
                                    9938414
                                                    101.
##
      1966 Afghanistan 1399999967.
                                   10152331
                                                    138.
  7
```

10372630

8 1967 Afghanistan 1673333418

59.8

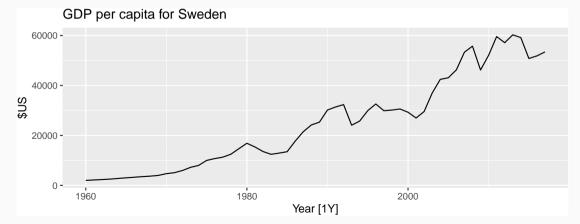
59.9

58.5

161

Data visualisation

```
gdppc %>%
filter(Country=="Sweden") %>%
autoplot(GDP_per_capita) +
   labs(title = "GDP per capita for Sweden", y = "$US")
```



Model estimation

The model() function trains models to data.

```
fit <- gdppc %>%
  model(trend_model = TSLM(GDP_per_capita ~ trend()))
fit
```

```
## # A mable: 263 x 2
```

```
## # Key: Country [263]
```

##		Country	trend_model
##		<fct></fct>	<model></model>
##	1	Afghanistan	<tslm></tslm>
##	2	Albania	<tslm></tslm>
##	3	Algeria	<tslm></tslm>
##	4	American Samoa	<tslm></tslm>
##	5	Andorra	<tslm></tslm>
##	6	Angola	<tslm></tslm>
##	7	Antigua and Barbuda	<tslm></tslm>

Model estimation

The model() function trains models to data.

```
fit <- gdppc %>%
  model(trend_model = TSLM(GDP_per_capita ~ trend()))
fit
```

```
## # A mable: 263 x 2
```

```
## # Key: Country [263]
```

##	Country	trend_model
##	<fct></fct>	<model></model>
##	1 Afghanistan	<tslm></tslm>
##	2 Albania	<tslm></tslm>
##	3 Algeria	<tslm></tslm>
##	4 American Samoa	<tslm></tslm>
##	5 Andorra	<tslm></tslm>
##	6 Angola	<tslm></tslm>
##	7 Antigua and Barbuda	<tslm></tslm>

A mable is a model table, each cell corresponds to a fitted model.

Producing forecasts

fit %>% forecast(h = "3 years")

```
## # A fable: 789 x 5 [1Y]
## # Key:
              Country, .model [263]
##
      Country
                     .model
                                  Year
                                         GDP_per_capita
                                                          .mean
##
      <fct>
                     <chr>
                                 <dbl>
                                                 <dist>
                                                          <dbl>
##
    1 Afghanistan
                   trend_model
                                  2018
                                           N(526, 9653)
                                                          526.
##
   2 Afghanistan
                    trend model
                                  2019
                                           N(534, 9689)
                                                          534.
   3 Afghanistan
                    trend_model
                                  2020
                                           N(542, 9727)
                                                          542.
##
##
   4 Albania
                     trend model
                                  2018
                                        N(4716, 476419)
                                                         4716.
##
   5 Albania
                     trend model
                                  2019
                                        N(4867, 481086)
                                                         4867.
   6 Albania
                     trend_model
##
                                  2020
                                        N(5018, 486012)
                                                          5018.
##
   7 Algeria
                     trend_model
                                  2018
                                        N(4410, 643094)
                                                         4410.
##
   8 Algeria
                     trend model
                                  2019
                                        N(4489, 645311)
                                                         4489.
   9 Algeria
##
                     trend model
                                  2020
                                        N(4568, 647602)
                                                         4568.
## 10 American Samoa trend model
                                  2018 N(12491, 652926) 12491.
```

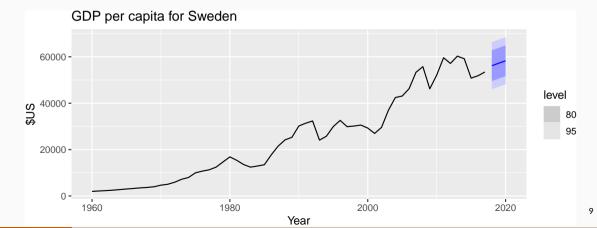
Producing forecasts

fit %>% forecast(h = "3 years")

		A fable: 789 x s Key: Country	5 [1Y] /, .model [20	A fable is a forecast table with point forecasts and distributions.			
##		Country	.model	Year	GDP_per_capica	·liean	
##		<fct></fct>	<chr></chr>	<dbl></dbl>	<dist></dist>	<dbl></dbl>	
##	1	Afghanistan	trend_model	2018	N(526, 9653)	526.	
##	2	Afghanistan	trend_model	2019	N(534, 9689)	534.	
##	3	Afghanistan	trend_model	2020	N(542, 9727)	542.	
##	4	Albania	trend_model	2018	N(4716, 476419)	4716.	
##	5	Albania	trend_model	2019	N(4867, 481086)	4867.	
##	6	Albania	trend_model	2020	N(5018, 486012)	5018.	
##	7	Algeria	trend_model	2018	N(4410, 643094)	4410.	
##	8	Algeria	trend_model	2019	N(4489, 645311)	4489.	
##	9	Algeria	trend_model	2020	N(4568, 647602)	4568.	
##	10	American Samoa	trend_model	2018	N(12491, 652926)	12491.	

Visualising forecasts

```
fit %>% forecast(h = "3 years") %>%
filter(Country=="Sweden") %>%
autoplot(gdppc) +
labs(title = "GDP per capita for Sweden", y = "$US")
```



Outline

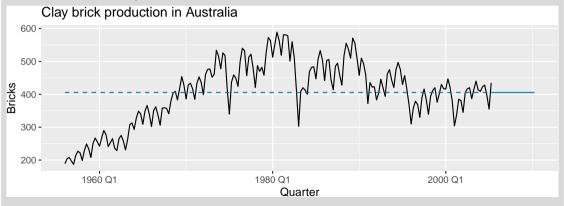
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MEAN(y): Average method

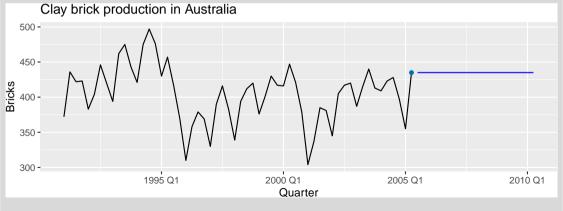
Forecast of all future values is equal to mean of historical data $\{y_1, \ldots, y_T\}$.

Forecasts:
$$\hat{y}_{T+h|T} = \bar{y} = (y_1 + \cdots + y_T)/T$$



NAIVE(y): Naïve method

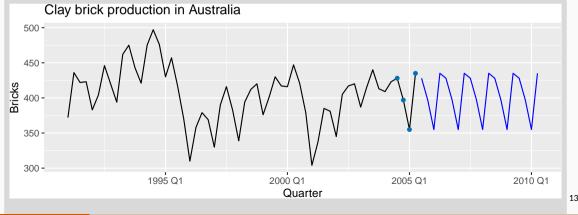
- Forecasts equal to last observed value.
- Forecasts: $\hat{y}_{T+h|T} = y_T$.
- Consequence of efficient market hypothesis.



SNAIVE(y ~ lag(m)): Seasonal naïve method

Forecasts equal to last value from same season.

Forecasts: $\hat{y}_{T+h|T} = y_{T+h-m(k+1)}$, where m = seasonal period and k is the integer part of (h - 1)/m.



RW(y ~ drift()): Drift method

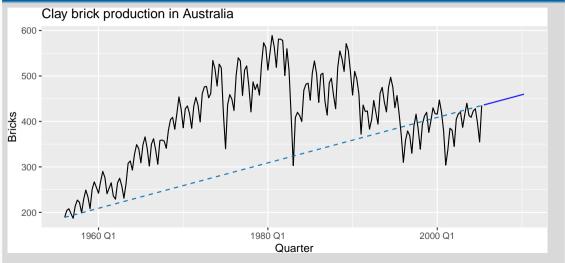
Forecasts equal to last value plus average change.

Forecasts:

$$\hat{y}_{T+h|T} = y_T + \frac{h}{T-1} \sum_{t=2}^{T} (y_t - y_{t-1})$$
$$= y_T + \frac{h}{T-1} (y_T - y_1).$$

Equivalent to extrapolating a line drawn between first and last observations.

Drift method



Model fitting

The model() function trains models to data.

```
brick_fit <- aus_production %>%
filter(!is.na(Bricks)) %>%
model(
   Seasonal_naive = SNAIVE(Bricks),
   Naive = NAIVE(Bricks),
   Drift = RW(Bricks ~ drift()),
   Mean = MEAN(Bricks)
)
```

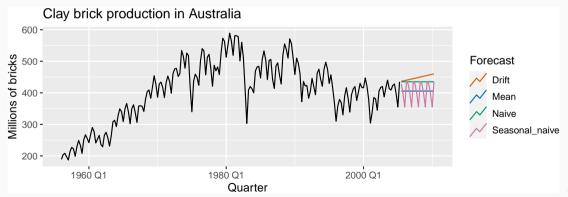
```
## # A mable: 1 x 4
## Seasonal_naive Naive Drift Mean
## <model> <model> <model> <model> <model>
## 1 <SNAIVE> <NAIVE> <RW w/ drift> <MEAN>
```

```
brick_fc <- brick_fit %>%
forecast(h = "5 years")
```

```
## # A fable: 80 x 4 [1Q]
## # Key: .model [4]
##
    .model
                  Quarter
                                Bricks .mean
##
   <chr>
                     <qtr> <dist> <dbl>
## 1 Seasonal_naive 2005 Q3 N(428, 2336)
                                         428
## 2 Seasonal_naive 2005 Q4 N(397, 2336)
                                         397
## 3 Seasonal_naive 2006 Q1 N(355, 2336)
                                         355
## 4 Seasonal_naive 2006 Q2 N(435, 2336)
                                         435
## # ... with 76 more rows
```

Visualising forecasts

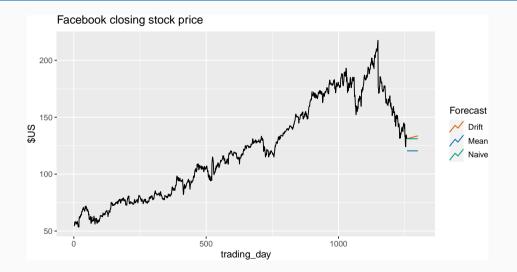
```
brick_fc %>%
  autoplot(aus_production, level = NULL) +
  labs(title = "Clay brick production in Australia",
        y = "Millions of bricks") +
   guides(colour = guide_legend(title = "Forecast"))
```



Extract training data

```
fb_stock <- gafa_stock %>%
  filter(Symbol == "FB") %>%
  mutate(trading_day = row_number()) %>%
  update_tsibble(index=trading_day, regular=TRUE)
```

```
# Specify, estimate and forecast
fb_stock %>%
  model(
    Mean = MEAN(Close),
    Naive = NAIVE(Close),
    Drift = RW(Close ~ drift())
) %>%
  forecast(h=42) %>%
  autoplot(fb_stock, level = NULL) +
  labs(title = "Facebook closing stock price", y="$US") +
  guides(colour=guide_legend(title="Forecast"))
```



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- $\hat{y}_{t|t-1}$ is the forecast of y_t based on observations y_1, \ldots, y_{t-1} .
- We call these "fitted values".
- Sometimes drop the subscript: $\hat{y}_t \equiv \hat{y}_{t|t-1}$.
- Often not true forecasts since parameters are estimated on all data.

For example:

- $\hat{y}_t = \bar{y}$ for average method.
- $\hat{y}_t = y_{t-1} + (y_T y_1)/(T 1)$ for drift method.

Residuals in forecasting: difference between observed value and its fitted value: $e_t = y_t - \hat{y}_{t|t-1}$.

Residuals in forecasting: difference between observed value and its fitted value: $e_t = y_t - \hat{y}_{t|t-1}$.

Assumptions

- 1 $\{e_t\}$ uncorrelated. If they aren't, then information left in residuals that should be used in computing forecasts.
- [2] { e_t } have mean zero. If they don't, then forecasts are biased.

Residuals in forecasting: difference between observed value and its fitted value: $e_t = y_t - \hat{y}_{t|t-1}$.

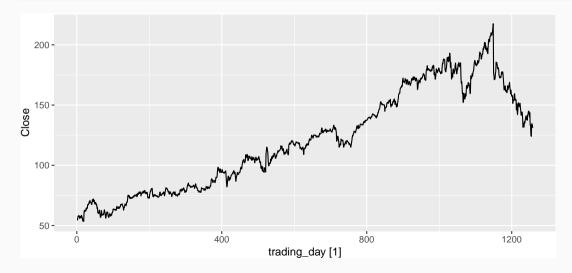
Assumptions

- 1 $\{e_t\}$ uncorrelated. If they aren't, then information left in residuals that should be used in computing forecasts.
- [2] { e_t } have mean zero. If they don't, then forecasts are biased.

Useful properties (for distributions & prediction intervals)

- [3] { e_t } have constant variance.
- ${}_{4}$ { e_t } are normally distributed.

fb_stock %>% autoplot(Close)



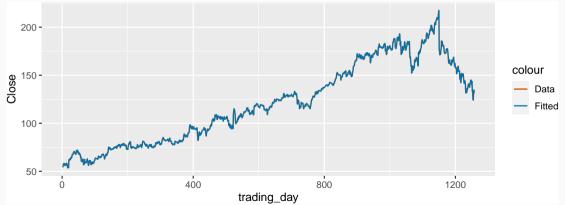
fit <- fb_stock %>% model(NAIVE(Close))
augment(fit)

##	# /	A tsibb]	le: 1,258 x 7	[1]				
##	# I	Key:	Symbol, .r	model [1]				
##		Symbol	.model	trading_day	Close	.fitted	.resid	.innov
##		<chr></chr>	<chr></chr>	<int></int>	<dbl></dbl>	<dbl></dbl>	<dbl></dbl>	<dbl></dbl>
##	1	FB	NAIVE(Close)	1	54.7	NA	NA	NA
##	2	FB	NAIVE(Close)	2	54.6	54.7	-0.150	-0.150
##	3	FB	NAIVE(Close)	3	57.2	54.6	2.64	2.64
##	4	FB	NAIVE(Close)	4	57.9	57.2	0.720	0.720
##	5	FB	NAIVE(Close)	5	58.2	57.9	0.310	0.310
##	6	FB	NAIVE(Close)	6	57.2	58.2	-1.01	-1.01
##	7	FB	NAIVE(Close)	7	57.9	57.2	0.720	0.720
##	8	FB	NAIVE(Close)	8	55.9	57.9	-2.03	-2.03
##	9	FB	NAIVE(Close)	9	57.7	55.9	1.83	1.83

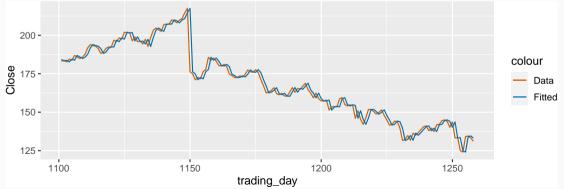
fit <- fb_stock %>% model(NAIVE(Close))
augment(fit)

## # A tsibble: 1,258 x 7 [1]									
##	# I	Key:	Symbol, .	model	[1]		$\hat{\mathbf{y}}_{t t-1}$	e_t	
##		Symbol	Symbol, .n .model	trad	ing_day	Close	.trueu	.resiu	.innov
##		<chr></chr>	<chr></chr>		<int></int>	<dbl></dbl>	<dbl></dbl>	<dbl></dbl>	<dbl></dbl>
##	1	FB	NAIVE(Close)		1	54.7	NA	NA	NA
##	2	FB	NAIVE(Close)		2	54.6	54.7	-0.150	-0.150
##	3	FB	NAIVE(Close)		3	57.2	54.6	2.64	2.64
##	4	FB	NAIVE(Close)		4	57.9	57.2	0.720	0.720
Naïve forecasts:			5	58.2	57.9	0.310	0.310		
					6	57.2	58.2	-1.01	-1.01
$\hat{\mathbf{y}}_{t t-1} = \mathbf{y}_{t-1}$			7	57.9	57.2	0.720	0.720		
	0.	=	$\hat{y}_{t t-1} = y_t -$	_ V.	8	55.9	57.9	-2.03	-2.03
	et	- yt -	$\mathbf{y}_{t t-1} - \mathbf{y}_{t}$	y t_	1 9	57.7	55.9	1.83	1.83

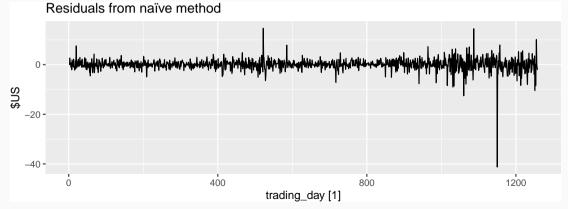
```
augment(fit) %>%
ggplot(aes(x = trading_day)) +
geom_line(aes(y = Close, colour = "Data")) +
geom_line(aes(y = .fitted, colour = "Fitted"))
```



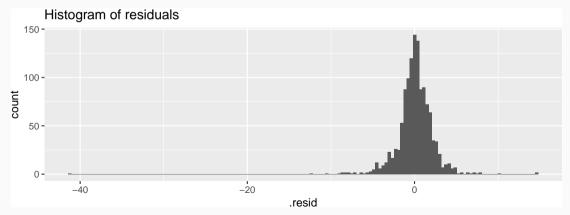
```
augment(fit) %>%
filter(trading_day > 1100) %>%
ggplot(aes(x = trading_day)) +
geom_line(aes(y = Close, colour = "Data")) +
geom_line(aes(y = .fitted, colour = "Fitted"))
```



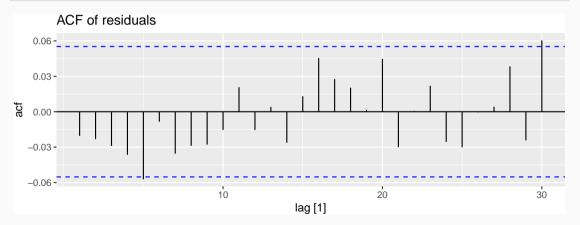
```
augment(fit) %>%
autoplot(.resid) +
labs(y = "$US",
    title = "Residuals from naïve method")
```



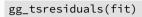
```
augment(fit) %>%
ggplot(aes(x = .resid)) +
geom_histogram(bins = 150) +
labs(title = "Histogram of residuals")
```

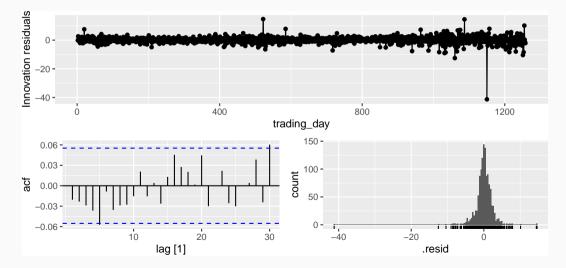


```
augment(fit) %>%
ACF(.resid) %>%
autoplot() + labs(title = "ACF of residuals")
```



gg_tsresiduals() function





- We assume that the residuals are white noise (uncorrelated, mean zero, constant variance). If they aren't, then there is information left in the residuals that should be used in computing forecasts.
- So a standard residual diagnostic is to check the ACF of the residuals of a forecasting method.
- We *expect* these to look like white noise.

Portmanteau tests

 r_k = autocorrelation of residual at lag k

Consider a whole set of r_k values, and develop a test to see whether the set is significantly different from a zero set.

Portmanteau tests

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Consider a whole set of r_k values, and develop a test to see whether the set is significantly different from a zero set.

Box-Pierce test

$$Q = T \sum_{k=1}^{\ell} r_k^2$$

where ℓ is max lag being considered and T is number of observations.

- If each r_k close to zero, Q will be **small**.
- If some r_k values large (positive or negative), Q will be **large**.

Portmanteau tests

r_k = autocorrelation of residual at lag k

Consider a whole set of r_k values, and develop a test to see whether the set is significantly different from a zero set.

Ljung-Box test

$$Q^* = T(T+2) \sum_{k=1}^{\ell} (T-k)^{-1} r_k^2$$

where ℓ is max lag being considered and T is number of observations.

- My preferences: *ℓ* = 10 for non-seasonal data, *h* = 2*m* for seasonal data (where *m* is seasonal period).
- Better performance, especially in small samples.

- If data are WN, Q^* has χ^2 distribution with (ℓK) degrees of freedom where K = no. parameters in model.
- When applied to raw data, set *K* = 0.
- lag = ℓ , dof = K

```
augment(fit) %>%
features(.resid, ljung_box, lag=10, dof=0)
```

```
## # A tibble: 1 x 4
## Symbol .model lb_stat lb_pvalue
## <chr> <chr> <chr> <dbl> <dbl> <dbl>
## 1 FB NAIVE(Close) 12.1 0.276
```

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- A forecast $\hat{y}_{T+h|T}$ is (usually) the mean of the conditional distribution $y_{T+h} \mid y_1, \ldots, y_T$.
- Most time series models produce normally distributed forecasts.
- The forecast distribution describes the probability of observing any future value.

Assuming residuals are normal, uncorrelated, sd = $\hat{\sigma}$:

Mean:	$y_{T+h T} \sim N(ar{y},(1+1/T)\hat{\sigma}^2)$
Naïve:	$\mathbf{y}_{T+h T}\sim N(\mathbf{y}_{T},\mathbf{h}\hat{\sigma}^2)$
Seasonal naïve:	$\mathbf{y}_{T+h T} \sim N(\mathbf{y}_{T+h-m(k+1)}, (k+1)\hat{\sigma}^2)$
Drift:	$\mathbf{y}_{T+h T} \sim N(\mathbf{y}_T + rac{h}{T-1}(\mathbf{y}_T - \mathbf{y}_1), hrac{T+h}{T}\hat{\sigma}^2)$

where k is the integer part of (h - 1)/m.

Note that when h = 1 and T is large, these all give the same approximate forecast variance: $\hat{\sigma}^2$.

- A prediction interval gives a region within which we expect y_{T+h} to lie with a specified probability.
- Assuming forecast errors are normally distributed, then a 95% PI is

$\hat{\mathbf{y}}_{\mathsf{T+h}|\mathsf{T}} \pm \mathbf{1.96} \hat{\sigma}_{h}$

where $\hat{\sigma}_h$ is the st dev of the *h*-step distribution.

• When h = 1, $\hat{\sigma}_h$ can be estimated from the residuals.

brick_fc %>% hilo(level = 95)

##	# A	A tsibble: 80 x	5 [10	5]					
##	# ł	Key: .mode	el [4]]					
##		.model	Quart	ter	I	Bricks	.mean		' 95% '
##		<chr></chr>	<qt< td=""><td>tr></td><td></td><td><dist></dist></td><td><dbl></dbl></td><td></td><td><hilo></hilo></td></qt<>	tr>		<dist></dist>	<dbl></dbl>		<hilo></hilo>
##	1	Seasonal_naive	2005	QЗ	N(428,	2336)	428	[333,	523]95
##	2	Seasonal_naive	2005	Q4	N(397,	2336)	397	[302,	492]95
##	3	Seasonal_naive	2006	Q1	N(355,	2336)	355	[260,	450]95
##	4	Seasonal_naive	2006	Q2	N(435,	2336)	435	[340,	530]95
##	5	Seasonal_naive	2006	QЗ	N(428,	4672)	428	[294,	562]95
##	6	Seasonal_naive	2006	Q4	N(397,	4672)	397	[263,	531]95
##	7	Seasonal_naive	2007	Q1	N(355,	4672)	355	[221,	489]95
##	8	Seasonal_naive	2007	Q2	N(435,	4672)	435	[301,	569]95
##	9	Seasonal_naive	2007	Q3	N(428,	7008)	428	[264,	592]95
##	10	Seasonal_naive	2007	Q4	N(397,	7008)	397	[233,	561]95

- Point forecasts often useless without a measure of uncertainty (such as prediction intervals).
- Prediction intervals require a stochastic model (with random errors, etc).
- For most models, prediction intervals get wider as the forecast horizon increases.
- Use level argument to control coverage.
- Check residual assumptions before believing them.
- Usually too narrow due to unaccounted uncertainty.

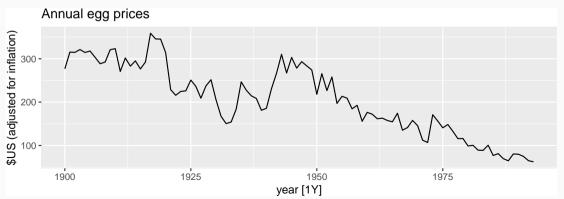
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Modelling with transformations

```
eggs <- prices %>%
filter(!is.na(eggs)) %>% select(eggs)
eggs %>% autoplot() +
labs(title="Annual egg prices",
    y="$US (adjusted for inflation)")
```



Transformations used in the left of the formula will be automatically back-transformed. To model log-transformed egg prices, you could use:

```
fit <- eggs %>%
  model(RW(log(eggs) ~ drift()))
fit
```

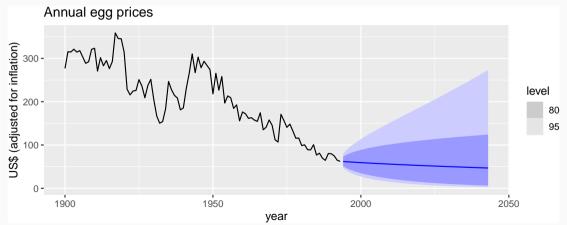
Forecasting with transformations

```
fc <- fit %>%
  forecast(h = 50)
fc
```

##	# A	A fable: 5	50 x	4	[1Y]				
##	# ł	Key: .	mode	ι	[1]				
##		.model				year		eggs	.mean
##		<chr></chr>				<dbl></dbl>		<dist></dist>	<dbl></dbl>
##	1	RW(log(eg	ggs)	~	drift())	1994	t(N(4.1,	0.018))	61.8
##	2	RW(log(eg	ggs)	~	drift())	1995	t(N(4.1,	0.036))	61.4
##	3	RW(log(eg	ggs)	~	drift())	1996	t(N(4.1,	0.054))	61.0
##	4	RW(log(eg	ggs)	~	drift())	1997	t(N(4.1,	0.073))	60.5
##	5	RW(log(eg	ggs)	~	drift())	1998	t(N(4.1,	0.093))	60.1
##	6	RW(log(eg	ggs)	~	drift())	1999	t(N(4,	, 0.11))	59.7
##	7	RW(log(eg	ggs)	~	drift())	2000	t(N(4,	, 0.13))	59.3
##	8	RW(log(eg	ggs)	~	drift())	2001	t(N(4,	, 0.15))	58.9
##	٥			~	drift())	ວ໑໑ວ	+ (N(A	0 17))	E0 C

Forecasting with transformations

```
fc %>% autoplot(eggs) +
   labs(title="Annual egg prices",
        y="US$ (adjusted for inflation)")
```



- Back-transformed point forecasts are medians.
- Back-transformed PI have the correct coverage.

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Back-transformed means

Let X be have mean μ and variance σ^2 .

Let f(x) be back-transformation function, and Y = f(X).

Taylor series expansion about μ :

$$f(X) = f(\mu) + (X - \mu)f'(\mu) + \frac{1}{2}(X - \mu)^2 f''(\mu).$$

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 $E[Y] = E[f(X)] = f(\mu) + \frac{1}{2}\sigma^2 f''(\mu)$

Bias adjustment

Box-Cox back-transformation:

$$y_t = \begin{cases} \exp(w_t) & \lambda = 0; \\ (\lambda W_t + 1)^{1/\lambda} & \lambda \neq 0. \end{cases}$$
$$f(x) = \begin{cases} e^x & \lambda = 0; \\ (\lambda x + 1)^{1/\lambda} & \lambda \neq 0. \end{cases}$$
$$f''(x) = \begin{cases} e^x & \lambda = 0; \\ (1 - \lambda)(\lambda x + 1)^{1/\lambda - 2} & \lambda \neq 0. \end{cases}$$

Bias adjustment

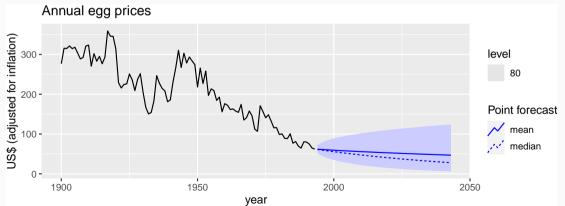
Box-Cox back-transformation:

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$$\mathsf{E}[\mathsf{Y}] = \begin{cases} e^{\mu} \left[1 + \frac{\sigma^2}{2} \right] & \lambda = 0; \\ (\lambda \mu + 1)^{1/\lambda} \left[1 + \frac{\sigma^2(1-\lambda)}{2(\lambda \mu + 1)^2} \right] & \lambda \neq 0. \end{cases}$$

Bias adjustment

```
fc %>%
  autoplot(eggs, level = 80, point_forecast = lst(mean, median)) +
  labs(title="Annual egg prices",
        y="US$ (adjusted for inflation)")
```



Outline

1 A tidy forecasting workflow

- 2 Some simple forecasting methods
- 3 Residual diagnostics
- 4 Distributional forecasts and prediction intervals
- 5 Forecasting with transformations
- 6 Forecasting and decomposition
- 7 Evaluating forecast accuracy

$$y_t = \hat{S}_t + \hat{A}_t$$

- Â_t is seasonally adjusted component
 Ŝ_t is seasonal component.
- Forecast \hat{S}_t using SNAIVE.
- Forecast \hat{A}_t using non-seasonal time series method.
- Combine forecasts of \hat{S}_t and \hat{A}_t to get forecasts of original data.

```
us_retail_employment <- us_employment %>%
  filter(year(Month) >= 1990, Title == "Retail Trade") %>%
  select(-Series_ID)
us_retail_employment
```

A tsibble: 357 x 3 [1M] Month Title Employed ## ## <mth> <chr> <dbl> ## 1 1990 Jan Retail Trade 13256. 2 1990 Feb Retail Trade ## 12966. 3 1990 Mär Retail Trade 12938. ## 4 1990 Apr Retail Trade 13012. ## 5 1990 Mai Retail Trade 13108. ## ## 6 1990 Jun Retail Trade 13183. ## 7 1990 Jul Retail Trade 13170. ## 8 1990 Aug Retail Trade 13160. ## 9 1990 Son Potail Trade 12112

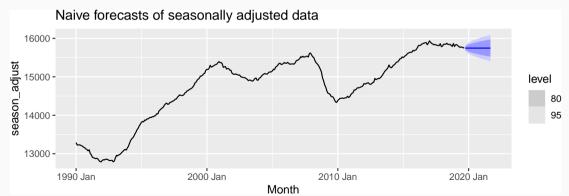
```
dcmp <- us_retail_employment %>%
  model(STL(Employed)) %>%
  components() %>% select(-.model)
dcmp
```

A tsibble: 357 x 6 [1M]

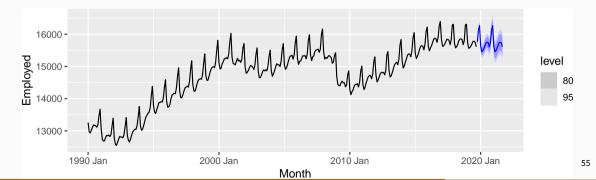
```
Month Employed trend season_year remainder season_adjust
##
##
       <mth>
               <dbl> <dbl>
                               <dbl>
                                         <dbl>
                                                     <dbl>
##
   1 1990 Jan 13256. 13288.
                              -33.0 0.836
                                                    13289.
   2 1990 Feb 12966. 13269.
##
                              -258.
                                       -44.6
                                                    13224.
##
   3 1990 Mär 12938. 13250.
                              -290.
                                       -22.1
                                                    13228.
   4 1990 Apr 13012. 13231.
                                       1.05
                                                    13232.
##
                              -220.
   5 1990 Mai 13108. 13211.
                                       11.3
                                                    13223.
##
                              -114.
##
   6 1990 Jun 13183. 13192. -24.3
                                       15.5
                                                    13207.
                                        21.6
##
   7 1990 Jul
              13170. 13172.
                               -23.2
                                                    13193.
   8 1990 Aug 13160. 13151. -9.52
                                        17.8
##
                                                    13169.
## 9 1990 Sen 13113 13131
                              -30 5
                                        22 0
                                                    12152
```

53

```
dcmp %>%
  model(NAIVE(season_adjust)) %>%
  forecast() %>%
  autoplot(dcmp) +
  labs(title = "Naive forecasts of seasonally adjusted data")
```



```
us_retail_employment %>%
model(stlf = decomposition_model(
   STL(Employed ~ trend(window = 7), robust = TRUE),
   NAIVE(season_adjust)
)) %>%
forecast() %>%
autoplot(us_retail_employment)
```



decomposition_model() creates a decomposition model

- You must provide a method for forecasting the season_adjust series.
- A seasonal naive method is used by default for the seasonal components.
- The variances from both the seasonally adjusted and seasonal forecasts are combined.

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- A model which fits the training data well will not necessarily forecast well.
- A perfect fit can always be obtained by using a model with enough parameters.
- Over-fitting a model to data is just as bad as failing to identify a systematic pattern in the data.
- The test set must not be used for any aspect of model development or calculation of forecasts.
- Forecast accuracy is based only on the test set.

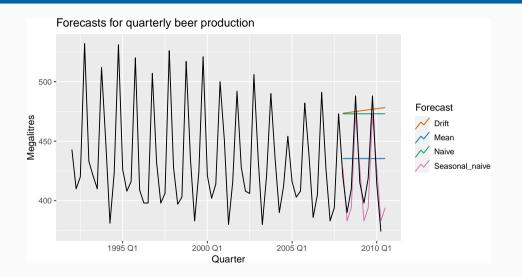
Forecast "error": the difference between an observed value and its forecast.

$$e_{T+h} = y_{T+h} - \hat{y}_{T+h|T},$$

where the training data is given by $\{y_1,\ldots,y_T\}$

- Unlike residuals, forecast errors on the test set involve multi-step forecasts.
- These are *true* forecast errors as the test data is not used in computing $\hat{y}_{T+h|T}$.

Measures of forecast accuracy



 $y_{T+h} = (T+h)$ th observation, h = 1, ..., H

 $\hat{y}_{T+h|T}$ = its forecast based on data up to time T.

 $e_{T+h} = y_{T+h} - \hat{y}_{T+h|T}$

MAE = mean($|e_{T+h}|$) MSE = mean(e_{T+h}^2) MAPE = 100mean($|e_{T+h}|/|y_{T+h}|$)

RMSE

 $=\sqrt{\mathrm{mean}(e_{T+h}^2)}$

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 $MAE = mean(|e_{T+h}|)$ $MSE = mean(e_{T+h}^2)$ $MAPE = 100mean(|e_{T+h}|/|y_{T+h}|)$ $RMSE = \sqrt{mean(e_{T+h}^2)}$

- MAE, MSE, RMSE are all scale dependent.
- MAPE is scale independent but is only sensible if $y_t \gg 0$ for all t, and y has a natural zero.

Measures of forecast accuracy

Mean Absolute Scaled Error

MASE = mean($|e_{T+h}|/Q$)

where Q is a stable measure of the scale of the time series $\{y_t\}$.

Proposed by Hyndman and Koehler (IJF, 2006).

For non-seasonal time series,

$$Q = (T-1)^{-1} \sum_{t=2}^{T} |y_t - y_{t-1}|$$

works well. Then MASE is equivalent to MAE relative to a naïve method.

Measures of forecast accuracy

Mean Absolute Scaled Error

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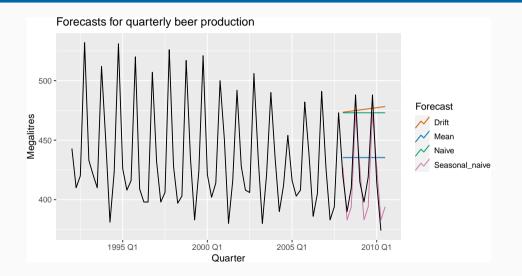
Proposed by Hyndman and Koehler (IJF, 2006).

For seasonal time series,

$$Q = (T - m)^{-1} \sum_{t=m+1}^{T} |y_t - y_{t-m}|$$

works well. Then MASE is equivalent to MAE relative to a seasonal naïve method.

Measures of forecast accuracy



```
recent_production <- aus_production %>%
  filter(vear(Quarter) >= 1992)
train <- recent production %>%
  filter(year(Quarter) <= 2007)</pre>
beer_fit <- train %>%
 model(
   Mean = MEAN(Beer),
   Naive = NAIVE(Beer),
    Seasonal_naive = SNAIVE(Beer),
    Drift = RW(Beer ~ drift())
beer fc <- beer fit %>%
  forecast(h = 10)
```

Measures of forecast accuracy

accuracy(beer_fit)

```
## # A tibble: 4 x 6
```

##		.model	.type	RMSE	MAE	MAPE	MASE	
##		<chr></chr>	<chr></chr>	<dbl></dbl>	<dbl></dbl>	<dbl></dbl>	<dbl></dbl>	
##	1	Drift	Training	65.3	54.8	12.2	3.83	
##	2	Mean	Training	43.6	35.2	7.89	2.46	
##	3	Naive	Training	65.3	54.7	12.2	3.83	
##	4	Seasonal_naive	Training	16.8	14.3	3.31	1	

accuracy(beer_fc, recent_production)

Traditional evaluation



Traditional evaluation



Time series cross-validation

h = 1_____ _____ ____ _ _____ _____ _____ ____ ____

Traditional evaluation



Time series cross-validation

h = 2_____ _____ _ ____ _ _____ _____ ____ ____

Traditional evaluation



Time series cross-validation

h = 3_____ _____ _____ ____ _____ ____ _ _____ ____ _____ ____

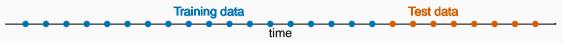
Traditional evaluation



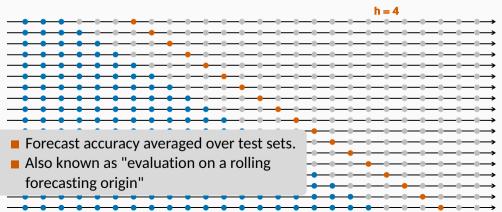
Time series cross-validation

h = 4_____ _____ ____ _ _____ ____ _____ ____ _ _____ _____ ____ _____ ____

Traditional evaluation



Time series cross-validation



Stretch with a minimum length of 3, growing by 1 each step.

```
fb_stretch <- fb_stock %>%
  stretch_tsibble(.init = 3, .step = 1) %>%
  filter(.id != max(.id))
```

##	#	A tsibble:	790,650 x 4 [1]	
##	#	Key:	.id [1,255]	
##		Date	Close trading_day .id	
##		<date></date>	<dbl> <int> <int></int></int></dbl>	
##	1	2014-01-02	54.7 1 1	
##	2	2014-01-03	54.6 2 1	
##	3	2014-01-06	57.2 3 1	
##	4	2014-01-02	54.7 1 2	
##	5	2014-01-03	54.6 2 2	
##	6	2014-01-06	57.2 3 2	
##	7	2014-01-07	57.9 4 2	

Estimate RW w/ drift models for each window.

```
fit_cv <- fb_stretch %>%
  model(RW(Close ~ drift()))
```

Produce one step ahead forecasts from all models.

fc_cv <- fit_cv %>%
forecast(h=1)

```
## # A fable: 1,255 x 5 [1]
## # Key: .id, Symbol [1,255]
## .id Symbol trading_day Close .mean
## <int> <chr> <dbl> <dist> <dbl><
dist> <dbl><
## 1 1 FB 4 N(58, 3.9) 58.4
## 2 2 FB 5 N(59, 2) 59.0
## 3 3 FB 6 N(59, 1.5) 59.1
## 4 4 FB 7 N(58, 1.8) 57.7
## # ... with 1,251 more rows</pre>
```

Cross-validated
fc_cv %>% accuracy(fb_stock)
Training set
fb_stock %>% model(RW(Close ~ drift())) %>% accuracy()

	RMSE	MAE	MAPE
Cross-validation	2.418	1.469	1.266
Training	2.414	1.465	1.261

A good way to choose the best forecasting model is to find the model with the smallest RMSE computed using time series cross-validation.