

Predictive Analytics

Ch8. Exponential smoothing

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- 1 Exponential smoothing
- 2 Simple exponential smoothing
- 3 Models with trend
- 4 Models with seasonality
- 5 Innovations state space models
- 6 Forecasting with exponential smoothing

- Developed in the 1950s and 1960s as methods (algorithms) to produce point forecasts.
- Combine a “level”, “trend” (slope) and “seasonal” component to describe a time series.
- The rate of change of the components are controlled by “smoothing parameters”: α , β and γ respectively.
- Need to choose best values for the smoothing parameters (and initial states).
- Equivalent ETS state space models developed in the 1990s and 2000s.

Big idea: control the rate of change

α controls the flexibility of the **level**

- If $\alpha = 0$, the level never updates (mean)
- If $\alpha = 1$, the level updates completely (naive)

β controls the flexibility of the **trend**

- If $\beta = 0$, the trend is linear
- If $\beta = 1$, the trend changes suddenly every observation

γ controls the flexibility of the **seasonality**

- If $\gamma = 0$, the seasonality is fixed (seasonal means)
- If $\gamma = 1$, the seasonality updates completely (seasonal naive)

A model for levels, trends, and seasonalities

We want a model that captures the level (ℓ_t), trend (b_t) and seasonality (s_t).

How do we combine these elements?

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Additively?

$$y_t = \ell_{t-1} + b_{t-1} + s_{t-m} + \varepsilon_t$$

Multiplicatively?

$$y_t = \ell_{t-1} b_{t-1} s_{t-m} (1 + \varepsilon_t)$$

Perhaps a mix of both?

$$y_t = (\ell_{t-1} + b_{t-1}) s_{t-m} + \varepsilon_t$$

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How do the level, trend and seasonal components evolve over time?

General notation

ETS : Exponential Smoothing
↑ ↑ ↑
Error Trend Season

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Error: Additive ("A") or multiplicative ("M")

Trend: None ("N"), additive ("A"), multiplicative ("M"), or damped ("Ad" or "Md").

Seasonality: None ("N"), additive ("A") or multiplicative ("M")

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Time series y_1, y_2, \dots, y_T .

Random walk forecasts

$$\hat{y}_{T+h|T} = y_T$$

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Random walk forecasts

$$\hat{y}_{T+h|T} = y_T$$

Average forecasts

$$\hat{y}_{T+h|T} = \frac{1}{T} \sum_{t=1}^T y_t$$

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Random walk forecasts

$$\hat{y}_{T+h|T} = y_T$$

Average forecasts

$$\hat{y}_{T+h|T} = \frac{1}{T} \sum_{t=1}^T y_t$$

- Want something in between these methods.
- Most recent data should have more weight.

Forecast equation

$$\hat{y}_{T+1|T} = \alpha y_T + \alpha(1 - \alpha)y_{T-1} + \alpha(1 - \alpha)^2 y_{T-2} + \dots,$$

where $0 \leq \alpha \leq 1$.

Forecast equation

$$\hat{y}_{T+1|T} = \alpha y_T + \alpha(1 - \alpha)y_{T-1} + \alpha(1 - \alpha)^2 y_{T-2} + \dots,$$

where $0 \leq \alpha \leq 1$.

Observation	Weights assigned to observations for:			
	$\alpha = 0.2$	$\alpha = 0.4$	$\alpha = 0.6$	$\alpha = 0.8$
y_T	0.2	0.4	0.6	0.8
y_{T-1}	0.16	0.24	0.24	0.16
y_{T-2}	0.128	0.144	0.096	0.032
y_{T-3}	0.1024	0.0864	0.0384	0.0064
y_{T-4}	$(0.2)(0.8)^4$	$(0.4)(0.6)^4$	$(0.6)(0.4)^4$	$(0.8)(0.2)^4$
y_{T-5}	$(0.2)(0.8)^5$	$(0.4)(0.6)^5$	$(0.6)(0.4)^5$	$(0.8)(0.2)^5$

Component form

Forecast equation

$$\hat{y}_{t+h|t} = l_t$$

Smoothing equation

$$l_t = \alpha y_t + (1 - \alpha)l_{t-1}$$

- l_t is the level (or the smoothed value) of the series at time t .
- $\hat{y}_{t+1|t} = \alpha y_t + (1 - \alpha)\hat{y}_{t|t-1}$

Simple Exponential Smoothing

Component form

Forecast equation

$$\hat{y}_{t+h|t} = l_t$$

Smoothing equation

$$l_t = \alpha y_t + (1 - \alpha)l_{t-1}$$

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- $\hat{y}_{t+1|t} = \alpha y_t + (1 - \alpha)\hat{y}_{t|t-1}$

Iterate to get exponentially weighted moving average form.

Weighted average form

$$\hat{y}_{T+1|T} = \sum_{j=0}^{T-1} \alpha(1 - \alpha)^j y_{T-j} + (1 - \alpha)^T l_0$$

- Need to choose best values for α and ℓ_0 .
- Similarly to regression, choose optimal parameters by minimising SSE:

$$\text{SSE} = \sum_{t=1}^T (y_t - \hat{y}_{t|t-1})^2.$$

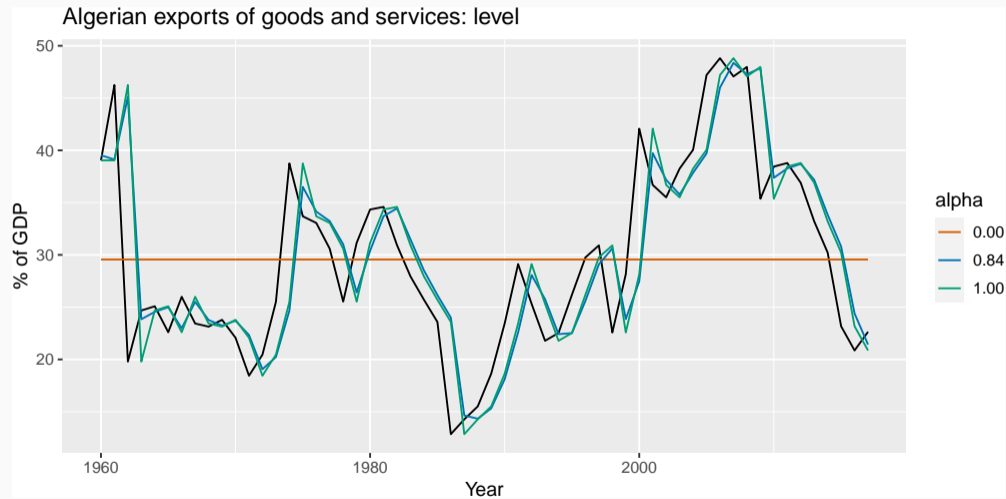
- Unlike regression there is no closed form solution — use numerical optimization.

- Need to choose best values for α and ℓ_0 .
- Similarly to regression, choose optimal parameters by minimising SSE:

$$\text{SSE} = \sum_{t=1}^T (y_t - \hat{y}_{t|t-1})^2.$$

- Unlike regression there is no closed form solution — use numerical optimization.
- For Algerian Exports example:
 - ▶ $\hat{\alpha} = 0.8400$
 - ▶ $\hat{\ell}_0 = 39.54$

Simple Exponential Smoothing



Methods

- Algorithms that return point forecasts.

Models

- Generate same point forecasts but can also generate forecast distributions.
- A stochastic (or random) data generating process that can generate an entire forecast distribution.
- Allow for “proper” model selection.

Component form

Forecast equation

$$\hat{y}_{t+h|t} = l_t$$

Smoothing equation

$$l_t = \alpha y_t + (1 - \alpha)l_{t-1}$$

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Smoothing equation

$$l_t = \alpha y_t + (1 - \alpha)l_{t-1}$$

Forecast error: $e_t = y_t - \hat{y}_{t|t-1} = y_t - l_{t-1}$.

Component form

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$$l_t = \alpha y_t + (1 - \alpha)l_{t-1}$$

Forecast error: $e_t = y_t - \hat{y}_{t|t-1} = y_t - l_{t-1}$.

Error correction form

$$y_t = l_{t-1} + e_t$$

$$l_t = l_{t-1} + \alpha(y_t - l_{t-1})$$

$$= l_{t-1} + \alpha e_t$$

Component form

Forecast equation

$$\hat{y}_{t+h|t} = l_t$$

Smoothing equation

$$l_t = \alpha y_t + (1 - \alpha)l_{t-1}$$

Forecast error: $e_t = y_t - \hat{y}_{t|t-1} = y_t - l_{t-1}$.

Error correction form

$$y_t = l_{t-1} + e_t$$

$$l_t = l_{t-1} + \alpha(y_t - l_{t-1})$$

$$= l_{t-1} + \alpha e_t$$

Specify probability distribution for e_t , we assume $e_t = \varepsilon_t \sim \text{NID}(0, \sigma^2)$.

Measurement equation

$$y_t = l_{t-1} + \varepsilon_t$$

State equation

$$l_t = l_{t-1} + \alpha \varepsilon_t$$

where $\varepsilon_t \sim \text{NID}(0, \sigma^2)$.

- “innovations” or “single source of error” because equations have the same error process, ε_t .
- Measurement equation: relationship between observations and states.
- State equation(s): evolution of the state(s) through time.

- Specify relative errors $\varepsilon_t = \frac{y_t - \hat{y}_{t|t-1}}{\hat{y}_{t|t-1}} \sim \text{NID}(0, \sigma^2)$
- Substituting $\hat{y}_{t|t-1} = l_{t-1}$ gives:
 - ▶ $y_t = l_{t-1} + l_{t-1}\varepsilon_t$
 - ▶ $e_t = y_t - \hat{y}_{t|t-1} = l_{t-1}\varepsilon_t$

- Specify relative errors $\varepsilon_t = \frac{y_t - \hat{y}_{t|t-1}}{\hat{y}_{t|t-1}} \sim \text{NID}(0, \sigma^2)$
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Measurement equation

$$y_t = l_{t-1}(1 + \varepsilon_t)$$

State equation

$$l_t = l_{t-1}(1 + \alpha\varepsilon_t)$$

ETS(M,N,N): SES with multiplicative errors.

- Specify relative errors $\varepsilon_t = \frac{y_t - \hat{y}_{t|t-1}}{\hat{y}_{t|t-1}} \sim \text{NID}(0, \sigma^2)$
- Substituting $\hat{y}_{t|t-1} = l_{t-1}$ gives:
 - ▶ $y_t = l_{t-1} + l_{t-1}\varepsilon_t$
 - ▶ $e_t = y_t - \hat{y}_{t|t-1} = l_{t-1}\varepsilon_t$

Measurement equation

$$y_t = l_{t-1}(1 + \varepsilon_t)$$

State equation

$$l_t = l_{t-1}(1 + \alpha\varepsilon_t)$$

- Models with additive and multiplicative errors with the same parameters generate the same point forecasts but different prediction intervals.

ETS(A,N,N): Specifying the model

```
ETS(y ~ error("A") + trend("N") + season("N"))
```

By default, an optimal value for α and l_0 is used.

α can be chosen manually in `trend()`.

```
trend("N", alpha = 0.5)  
trend("N", alpha_range = c(0.2, 0.8))
```


Example: Algerian Exports

```
algeria_economy <- global_economy %>%  
  filter(Country == "Algeria")  
fit <- algeria_economy %>%  
  model(ANN = ETS(Exports ~ error("A") + trend("N") + season("N")))  
report(fit)
```

```
## Series: Exports
```

```
## Model: ETS(A,N,N)
```

```
## Smoothing parameters:
```

```
## alpha = 0.84
```

```
##
```

```
## Initial states:
```

```
## l[0]
```

```
## 39.5
```

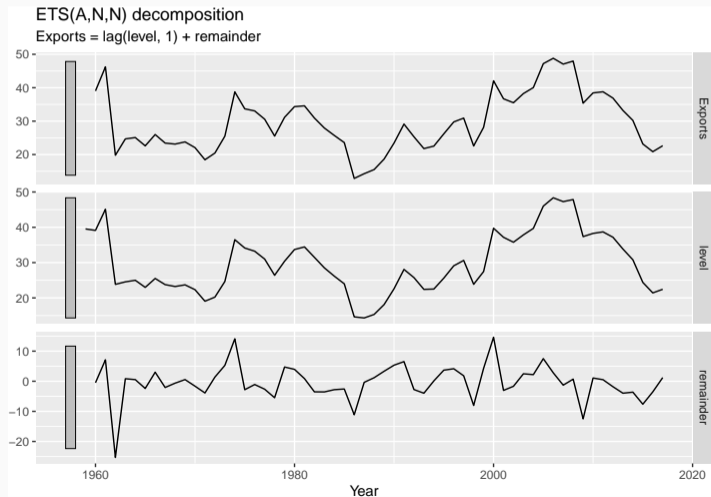
```
##
```

```
## sigma^2: 35.6
```

```
##
```

Example: Algerian Exports

```
components(fit) %>% autoplot()
```



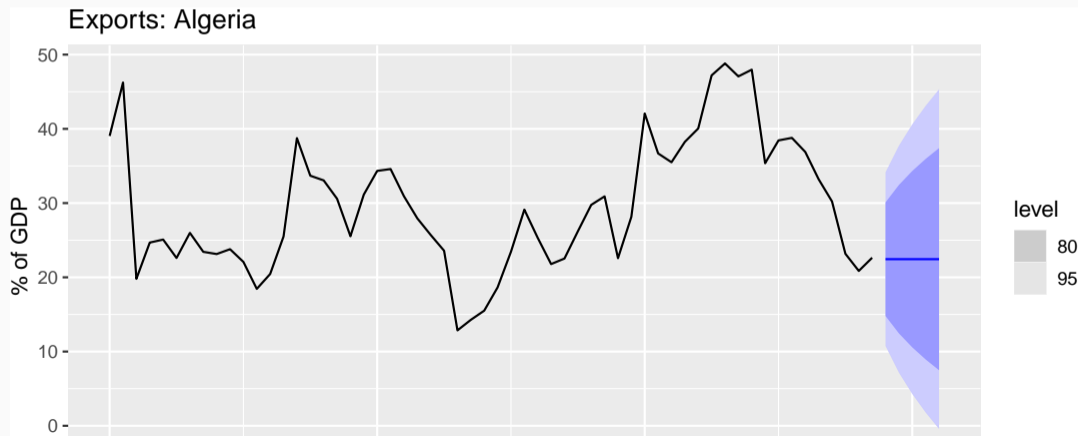
Example: Algerian Exports

```
components(fit) %>%  
  left_join(fitted(fit), by = c("Country", ".model", "Year"))
```

```
## # A dable: 59 x 7 [1Y]  
## # Key:      Country, .model [1]  
## # :        Exports = lag(level, 1) + remainder  
##   Country .model Year Exports level remainder .fitted  
##   <fct>   <chr> <dbl> <dbl> <dbl> <dbl> <dbl>  
## 1 Algeria ANN    1959  NA    39.5  NA     NA     NA  
## 2 Algeria ANN    1960  39.0  39.1  -0.496 39.5  
## 3 Algeria ANN    1961  46.2  45.1   7.12  39.1  
## 4 Algeria ANN    1962  19.8  23.8 -25.3  45.1  
## 5 Algeria ANN    1963  24.7  24.6   0.841 23.8  
## 6 Algeria ANN    1964  25.1  25.0   0.534 24.6  
## 7 Algeria ANN    1965  22.6  23.0  -2.39  25.0  
## 8 Algeria ANN    1966  26.0  25.5   3.00  23.0
```

Example: Algerian Exports

```
fit %>%  
  forecast(h = 5) %>%  
  autoplot(algeria_economy) +  
  labs(y = "% of GDP", title = "Exports: Algeria")
```



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Component form

Forecast

$$\hat{y}_{t+h|t} = l_t + hb_t$$

Level

$$l_t = \alpha y_t + (1 - \alpha)(l_{t-1} + b_{t-1})$$

Trend

$$b_t = \beta^*(l_t - l_{t-1}) + (1 - \beta^*)b_{t-1},$$

Component form

Forecast	$\hat{y}_{t+h t} = l_t + hb_t$
Level	$l_t = \alpha y_t + (1 - \alpha)(l_{t-1} + b_{t-1})$
Trend	$b_t = \beta^*(l_t - l_{t-1}) + (1 - \beta^*)b_{t-1},$

- Two smoothing parameters α and β^* ($0 \leq \alpha, \beta^* \leq 1$).
- l_t level: weighted average between y_t and one-step ahead forecast for time t , ($l_{t-1} + b_{t-1} = \hat{y}_{t|t-1}$)
- b_t slope: weighted average of $(l_t - l_{t-1})$ and b_{t-1} , current and previous estimate of slope.
- Choose $\alpha, \beta^*, l_0, b_0$ to minimise SSE.

Holt's linear method with additive errors.

- Assume $\varepsilon_t = y_t - l_{t-1} - b_{t-1} \sim \text{NID}(0, \sigma^2)$.
- Substituting into the error correction equations for Holt's linear method

$$y_t = l_{t-1} + b_{t-1} + \varepsilon_t$$

$$l_t = l_{t-1} + b_{t-1} + \alpha\varepsilon_t$$

$$b_t = b_{t-1} + \alpha\beta^*\varepsilon_t$$

- For simplicity, set $\beta = \alpha\beta^*$.

Holt's linear method with multiplicative errors.

- Assume $\varepsilon_t = \frac{y_t - (l_{t-1} + b_{t-1})}{(l_{t-1} + b_{t-1})}$
- Following a similar approach as above, the innovations state space model underlying Holt's linear method with multiplicative errors is specified as

$$y_t = (l_{t-1} + b_{t-1})(1 + \varepsilon_t)$$

$$l_t = (l_{t-1} + b_{t-1})(1 + \alpha\varepsilon_t)$$

$$b_t = b_{t-1} + \beta(l_{t-1} + b_{t-1})\varepsilon_t$$

where again $\beta = \alpha\beta^*$ and $\varepsilon_t \sim \text{NID}(0, \sigma^2)$.

ETS(A,A,N): Specifying the model

```
ETS(y ~ error("A") + trend("A") + season("N"))
```

By default, optimal values for β and b_0 are used.

β can be chosen manually in `trend()`.

```
trend("A", beta = 0.004)  
trend("A", beta_range = c(0, 0.1))
```

Example: Australian population

```
aus_economy <- global_economy %>% filter(Code == "AUS") %>%  
  mutate(Pop = Population / 1e6)  
fit <- aus_economy %>%  
  model(AAN = ETS(Pop ~ error("A") + trend("A") + season("N")))  
report(fit)
```

```
## Series: Pop
```

```
## Model: ETS(A,A,N)
```

```
## Smoothing parameters:
```

```
## alpha = 1
```

```
## beta = 0.327
```

```
##
```

```
## Initial states:
```

```
## l[0] b[0]
```

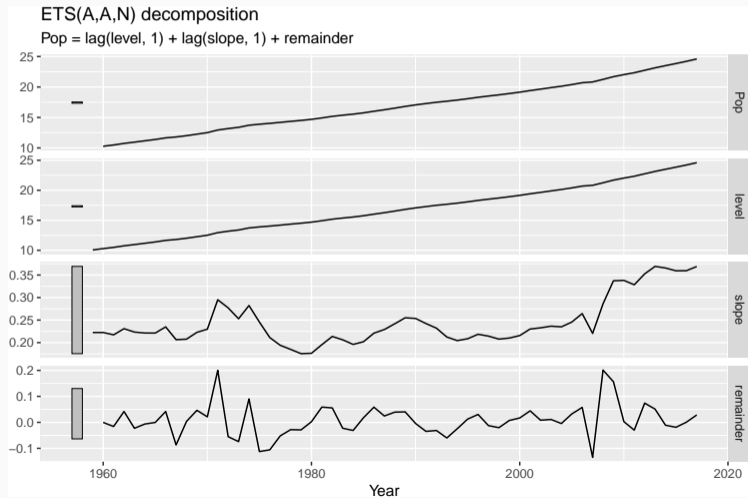
```
## 10.1 0.222
```

```
##
```

```
## sigma^2: 0.0041
```

Example: Australian population

```
components(fit) %>% autoplot()
```

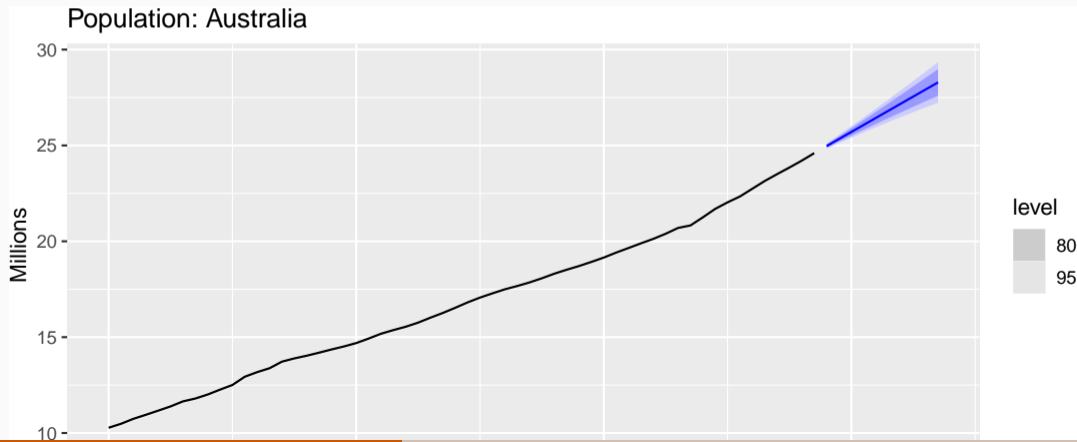


Example: Australian population

```
components(fit) %>%  
  left_join(fitted(fit), by = c("Country", ".model", "Year"))  
  
## # A dable: 59 x 8 [1Y]  
## # Key:      Country, .model [1]  
## # :        Pop = lag(level, 1) + lag(slope, 1) + remainder  
##   Country  .model  Year  Pop level slope remainder .fitted  
##   <fct>    <chr>  <dbl> <dbl> <dbl> <dbl>      <dbl>  <dbl>  
## 1 Australia AAN      1959  NA    10.1 0.222 NA          NA  
## 2 Australia AAN      1960  10.3  10.3 0.222 -0.000145  10.3  
## 3 Australia AAN      1961  10.5  10.5 0.217 -0.0159    10.5  
## 4 Australia AAN      1962  10.7  10.7 0.231  0.0418    10.7  
## 5 Australia AAN      1963  11.0  11.0 0.223 -0.0229    11.0  
## 6 Australia AAN      1964  11.2  11.2 0.221 -0.00641   11.2  
## 7 Australia AAN      1965  11.4  11.4 0.221 -0.000314  11.4  
## 8 Australia AAN      1966  11.7  11.7 0.235  0.0418    11.6
```

Example: Australian population

```
fit %>%  
  forecast(h = 10) %>%  
  autoplot(aus_economy) +  
  labs(y = "Millions", title = "Population: Australia")
```



Component form

$$\hat{y}_{t+h|t} = l_t + (\phi + \phi^2 + \dots + \phi^h)b_t$$

$$l_t = \alpha y_t + (1 - \alpha)(l_{t-1} + \phi b_{t-1})$$

$$b_t = \beta^*(l_t - l_{t-1}) + (1 - \beta^*)\phi b_{t-1}.$$

Component form

$$\hat{y}_{t+h|t} = l_t + (\phi + \phi^2 + \dots + \phi^h)b_t$$

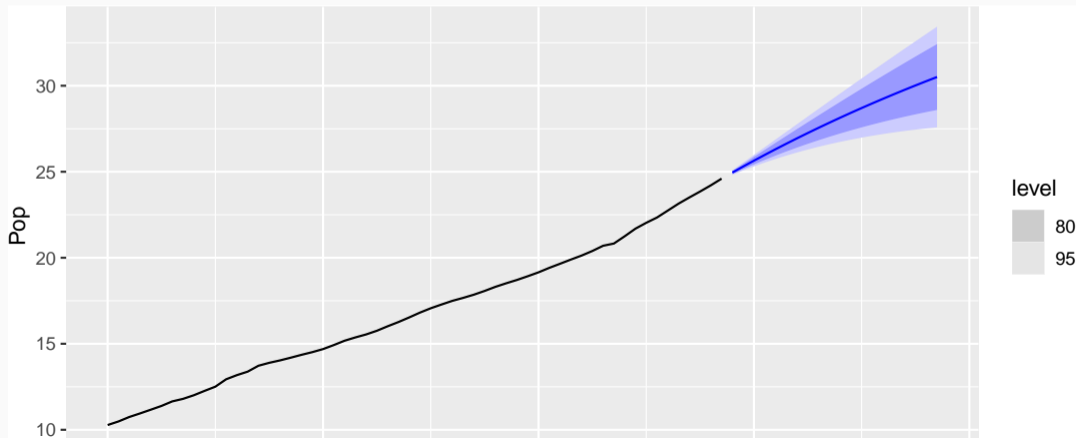
$$l_t = \alpha y_t + (1 - \alpha)(l_{t-1} + \phi b_{t-1})$$

$$b_t = \beta^*(l_t - l_{t-1}) + (1 - \beta^*)\phi b_{t-1}.$$

- Damping parameter $0 < \phi < 1$.
- If $\phi = 1$, identical to Holt's linear trend.
- As $h \rightarrow \infty$, $\hat{y}_{T+h|T} \rightarrow l_T + \phi b_T / (1 - \phi)$.
- Short-run forecasts trended, long-run forecasts constant.

Example: Australian population

```
aus_economy %>%  
  model(holt = ETS(Pop ~ error("A") + trend("Ad") + season("N"))) %>%  
  forecast(h = 20) %>%  
  autoplot(aus_economy)
```



Example: Australian population

```
fit <- aus_economy %>%  
  filter(Year <= 2010) %>%  
  model(  
    ses = ETS(Pop ~ error("A") + trend("N") + season("N")),  
    holt = ETS(Pop ~ error("A") + trend("A") + season("N")),  
    damped = ETS(Pop ~ error("A") + trend("Ad") + season("N"))  
  )
```

```
tidy(fit)  
accuracy(fit)
```

Example: Australian population

	term	SES	Linear trend	Damped trend
	α	1.00	1.00	1.00
	β^*		0.30	0.40
	ϕ			0.98
	NA		0.22	0.25
	NA	10.28	10.05	10.04
Training RMSE		0.24	0.06	0.07
Test RMSE		1.63	0.15	0.21
Test MASE		6.18	0.55	0.75
Test MAPE		6.09	0.55	0.74
Test MAE		1.45	0.13	0.18

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Holt and Winters extended Holt's method to capture seasonality.

Component form

$$\hat{y}_{t+h|t} = l_t + hb_t + s_{t+h-m(k+1)}$$

$$l_t = \alpha(y_t - s_{t-m}) + (1 - \alpha)(l_{t-1} + b_{t-1})$$

$$b_t = \beta^*(l_t - l_{t-1}) + (1 - \beta^*)b_{t-1}$$

$$s_t = \gamma(y_t - l_{t-1} - b_{t-1}) + (1 - \gamma)s_{t-m}$$

- $k = \text{integer part of } (h - 1)/m$. Ensures estimates from the final year are used for forecasting.
- Parameters: $0 \leq \alpha \leq 1$, $0 \leq \beta^* \leq 1$, $0 \leq \gamma \leq 1 - \alpha$ and $m = \text{period of seasonality}$ (e.g. $m = 4$ for quarterly data).

- Seasonal component is usually expressed as $s_t = \gamma^*(y_t - \ell_t) + (1 - \gamma^*)s_{t-m}$.
- Substitute in for ℓ_t : $s_t = \gamma^*(1 - \alpha)(y_t - \ell_{t-1} - b_{t-1}) + [1 - \gamma^*(1 - \alpha)]s_{t-m}$
- We set $\gamma = \gamma^*(1 - \alpha)$.
- The usual parameter restriction is $0 \leq \gamma^* \leq 1$, which translates to $0 \leq \gamma \leq (1 - \alpha)$.

Holt-Winters additive method with additive errors.

Forecast equation

$$\hat{y}_{t+h|t} = l_t + hb_t + s_{t+h-m(k+1)}$$

Observation equation

$$y_t = l_{t-1} + b_{t-1} + s_{t-m} + \varepsilon_t$$

State equations

$$l_t = l_{t-1} + b_{t-1} + \alpha\varepsilon_t$$

$$b_t = b_{t-1} + \beta\varepsilon_t$$

$$s_t = s_{t-m} + \gamma\varepsilon_t$$

- Forecast errors: $\varepsilon_t = y_t - \hat{y}_{t|t-1}$
- k is integer part of $(h - 1)/m$.

Seasonal variations change in proportion to the level of the series.

Component form

$$\hat{y}_{t+h|t} = (l_t + hb_t)s_{t+h-m(k+1)}$$

$$l_t = \alpha \frac{y_t}{s_{t-m}} + (1 - \alpha)(l_{t-1} + b_{t-1})$$

$$b_t = \beta^*(l_t - l_{t-1}) + (1 - \beta^*)b_{t-1}$$

$$s_t = \gamma \frac{y_t}{(l_{t-1} + b_{t-1})} + (1 - \gamma)s_{t-m}$$

- k is integer part of $(h - 1)/m$.
- Additive method: s_t in absolute terms — within each year $\sum_i s_i \approx 0$.
- Multiplicative method: s_t in relative terms — within each year $\sum_i s_i \approx m$.

Holt-Winters multiplicative method with multiplicative errors.

Forecast equation	$\hat{y}_{t+h t} = (l_t + hb_t)s_{t+h-m(k+1)}$
Observation equation	$y_t = (l_{t-1} + b_{t-1})s_{t-m}(1 + \varepsilon_t)$
State equations	$l_t = (l_{t-1} + b_{t-1})(1 + \alpha\varepsilon_t)$ $b_t = b_{t-1} + \beta(l_{t-1} + b_{t-1})\varepsilon_t$ $s_t = s_{t-m}(1 + \gamma\varepsilon_t)$

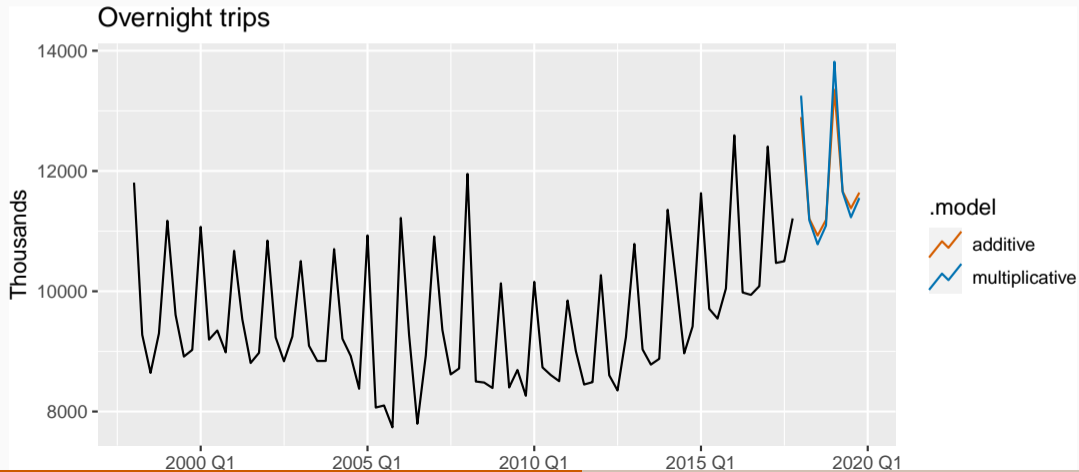
- Forecast errors: $\varepsilon_t = (y_t - \hat{y}_{t|t-1})/\hat{y}_{t|t-1}$
- k is integer part of $(h - 1)/m$.

Example: Australian holiday tourism

```
aus_holidays <- tourism %>%
  filter(Purpose == "Holiday") %>%
  summarise(Trips = sum(Trips))
fit <- aus_holidays %>%
  model(
    additive = ETS(Trips ~ error("A") + trend("A") + season("A")),
    multiplicative = ETS(Trips ~ error("M") + trend("A") + season("M"))
  )
fc <- fit %>% forecast()
```

Example: Australian holiday tourism

```
fc %>%  
  autoplot(aus_holidays, level = NULL) +  
  labs(y = "Thousands", title = "Overnight trips")
```



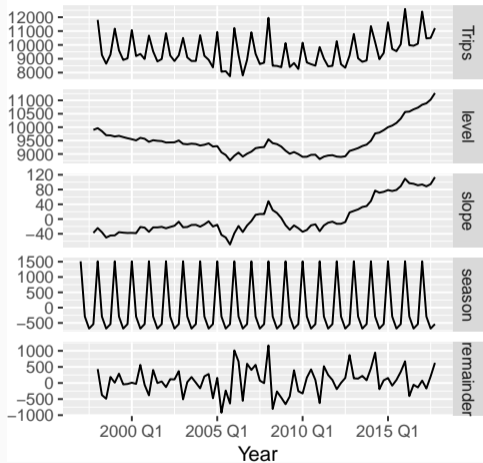
Estimated components

```
components(fit)
```

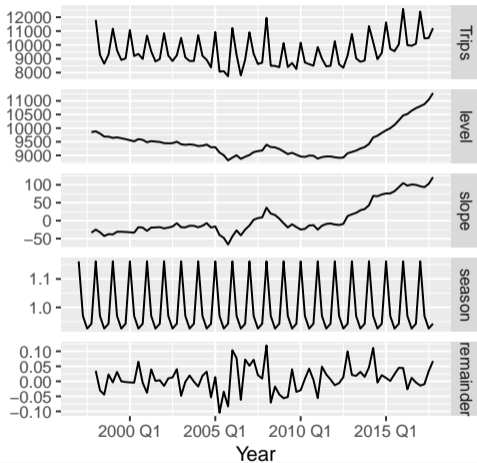
```
## # A dable: 168 x 7 [1Q]
## # Key:      .model [2]
## # :        Trips = lag(level, 1) + lag(slope, 1) + lag(season, 4) +
## # remainder
##   .model   Quarter   Trips level slope season remainder
##   <chr>    <qtr>    <dbl> <dbl> <dbl> <dbl>    <dbl>
## 1 additive 1997 Q1      NA     NA    NA    1512.     NA
## 2 additive 1997 Q2      NA     NA    NA    -290.     NA
## 3 additive 1997 Q3      NA     NA    NA    -684.     NA
## 4 additive 1997 Q4      NA  9899.  -37.4  -538.     NA
## 5 additive 1998 Q1 11806.  9964.  -24.5  1512.    433.
## 6 additive 1998 Q2  9276.  9851.  -35.6  -290.   -374.
## 7 additive 1998 Q3  8642.  9700.  -50.2  -684.   -489.
## 8 additive 1998 Q4  9300.  9694.  -44.6  -538.    188.
```

Estimated components

Additive states



Multiplicative states



Often the single most accurate forecasting method for seasonal data:

$$\hat{y}_{t+h|t} = [l_t + (\phi + \phi^2 + \dots + \phi^h)b_t]s_{t+h-m(k+1)}$$

$$l_t = \alpha(y_t/s_{t-m}) + (1 - \alpha)(l_{t-1} + \phi b_{t-1})$$

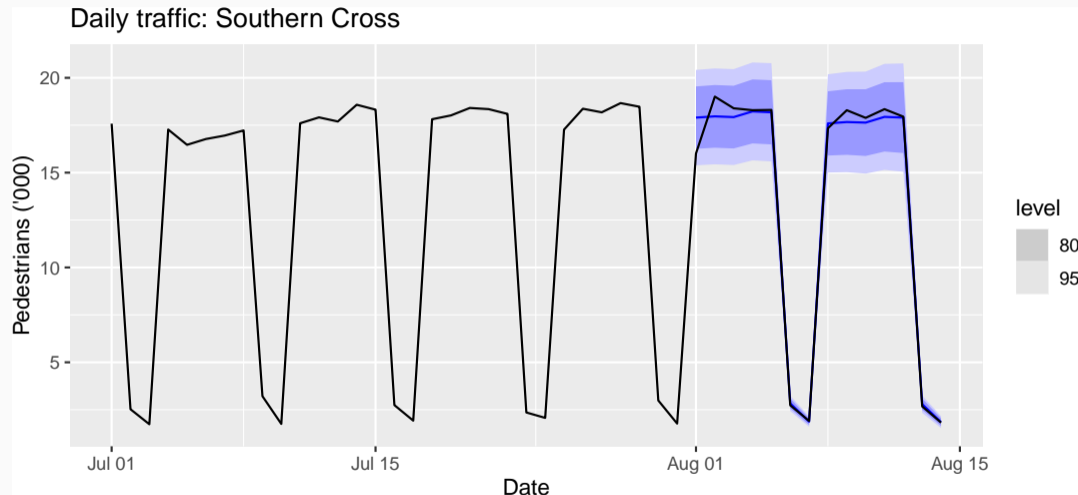
$$b_t = \beta^*(l_t - l_{t-1}) + (1 - \beta^*)\phi b_{t-1}$$

$$s_t = \gamma \frac{y_t}{(l_{t-1} + \phi b_{t-1})} + (1 - \gamma)s_{t-m}$$

Holt-Winters with daily data

```
sth_cross_ped <- pedestrian %>%
  filter(
    Date >= "2016-07-01",
    Sensor == "Southern Cross Station"
  ) %>%
  index_by(Date) %>%
  summarise(Count = sum(Count) / 1000)
sth_cross_ped %>%
  filter(Date <= "2016-07-31") %>%
  model(
    hw = ETS(Count ~ error("M") + trend("Ad") + season("M"))
  ) %>%
  forecast(h = "2 weeks") %>%
  autoplot(sth_cross_ped %>% filter(Date <= "2016-08-14")) +
  labs(
    title = "Daily traffic: Southern Cross",
    y = "Pedestrians ('000)"
  )
```

Holt-Winters with daily data



- 1 Exponential smoothing
- 2 Simple exponential smoothing
- 3 Models with trend
- 4 Models with seasonality
- 5 Innovations state space models**
- 6 Forecasting with exponential smoothing

Exponential smoothing methods

Trend Component		Seasonal Component		
		N (None)	A (Additive)	M (Multiplicative)
N	(None)	(N,N)	(N,A)	(N,M)
A	(Additive)	(A,N)	(A,A)	(A,M)
A _d	(Additive damped)	(A _d ,N)	(A _d ,A)	(A _d ,M)

(N,N): Simple exponential smoothing

(A,N): Holt's linear method

(A_d,N): Additive damped trend method

(A,A): Additive Holt-Winters' method

(A,M): Multiplicative Holt-Winters' method

(A_d,M): Damped multiplicative Holt-Winters' method

Exponential smoothing methods

Trend Component		Seasonal Component		
		N (None)	A (Additive)	M (Multiplicative)
N	(None)	(N,N)	(N,A)	(N,M)
A	(Additive)	(A,N)	(A,A)	(A,M)
A_d	(Additive damped)	(A_d ,N)	(A_d ,A)	(A_d ,M)

(N,N): Simple exponential smoothing

(A,N): Holt's linear method

(A_d ,N): Additive damped trend method

(A,A): Additive Holt-Winters' method

(A,M): Multiplicative Holt-Winters' method

(A_d ,M): Damped multiplicative Holt-Winters' method

There are also multiplicative trend methods (not recommended).

Additive Error

Trend Component		Seasonal Component		
		N (None)	A (Additive)	M (Multiplicative)
N	(None)	A,N,N	A,N,A	A,N,M
A	(Additive)	A,A,N	A,A,A	A,A,M
A _d	(Additive damped)	A,A _d ,N	A,A _d ,A	A,A _d ,M

Multiplicative Error

Trend Component		Seasonal Component		
		N (None)	A (Additive)	M (Multiplicative)
N	(None)	M,N,N	M,N,A	M,N,M
A	(Additive)	M,A,N	M,A,A	M,A,M
A _d	(Additive damped)	M,A _d ,N	M,A _d ,A	M,A _d ,M

Additive error models

Trend	Seasonal		
	N	A	M
N	$y_t = \ell_{t-1} + \varepsilon_t$ $\ell_t = \ell_{t-1} + \alpha\varepsilon_t$	$y_t = \ell_{t-1} + s_{t-m} + \varepsilon_t$ $\ell_t = \ell_{t-1} + \alpha\varepsilon_t$ $s_t = s_{t-m} + \gamma\varepsilon_t$	$y_t = \ell_{t-1}s_{t-m} + \varepsilon_t$ $\ell_t = \ell_{t-1} + \alpha\varepsilon_t/s_{t-m}$ $s_t = s_{t-m} + \gamma\varepsilon_t/\ell_{t-1}$
A	$y_t = \ell_{t-1} + b_{t-1} + \varepsilon_t$ $\ell_t = \ell_{t-1} + b_{t-1} + \alpha\varepsilon_t$ $b_t = b_{t-1} + \beta\varepsilon_t$	$y_t = \ell_{t-1} + b_{t-1} + s_{t-m} + \varepsilon_t$ $\ell_t = \ell_{t-1} + b_{t-1} + \alpha\varepsilon_t$ $b_t = b_{t-1} + \beta\varepsilon_t$ $s_t = s_{t-m} + \gamma\varepsilon_t$	$y_t = (\ell_{t-1} + b_{t-1})s_{t-m} + \varepsilon_t$ $\ell_t = \ell_{t-1} + b_{t-1} + \alpha\varepsilon_t/s_{t-m}$ $b_t = b_{t-1} + \beta\varepsilon_t/s_{t-m}$ $s_t = s_{t-m} + \gamma\varepsilon_t/(\ell_{t-1} + b_{t-1})$
A_d	$y_t = \ell_{t-1} + \phi b_{t-1} + \varepsilon_t$ $\ell_t = \ell_{t-1} + \phi b_{t-1} + \alpha\varepsilon_t$ $b_t = \phi b_{t-1} + \beta\varepsilon_t$	$y_t = \ell_{t-1} + \phi b_{t-1} + s_{t-m} + \varepsilon_t$ $\ell_t = \ell_{t-1} + \phi b_{t-1} + \alpha\varepsilon_t$ $b_t = \phi b_{t-1} + \beta\varepsilon_t$ $s_t = s_{t-m} + \gamma\varepsilon_t$	$y_t = (\ell_{t-1} + \phi b_{t-1})s_{t-m} + \varepsilon_t$ $\ell_t = \ell_{t-1} + \phi b_{t-1} + \alpha\varepsilon_t/s_{t-m}$ $b_t = \phi b_{t-1} + \beta\varepsilon_t/s_{t-m}$ $s_t = s_{t-m} + \gamma\varepsilon_t/(\ell_{t-1} + \phi b_{t-1})$

Multiplicative error models

Trend	Seasonal		
	N	A	M
N	$y_t = \ell_{t-1}(1 + \varepsilon_t)$ $\ell_t = \ell_{t-1}(1 + \alpha\varepsilon_t)$	$y_t = (\ell_{t-1} + s_{t-m})(1 + \varepsilon_t)$ $\ell_t = \ell_{t-1} + \alpha(\ell_{t-1} + s_{t-m})\varepsilon_t$ $s_t = s_{t-m} + \gamma(\ell_{t-1} + s_{t-m})\varepsilon_t$	$y_t = \ell_{t-1}s_{t-m}(1 + \varepsilon_t)$ $\ell_t = \ell_{t-1}(1 + \alpha\varepsilon_t)$ $s_t = s_{t-m}(1 + \gamma\varepsilon_t)$
A	$y_t = (\ell_{t-1} + b_{t-1})(1 + \varepsilon_t)$ $\ell_t = (\ell_{t-1} + b_{t-1})(1 + \alpha\varepsilon_t)$ $b_t = b_{t-1} + \beta(\ell_{t-1} + b_{t-1})\varepsilon_t$	$y_t = (\ell_{t-1} + b_{t-1} + s_{t-m})(1 + \varepsilon_t)$ $\ell_t = \ell_{t-1} + b_{t-1} + \alpha(\ell_{t-1} + b_{t-1} + s_{t-m})\varepsilon_t$ $b_t = b_{t-1} + \beta(\ell_{t-1} + b_{t-1} + s_{t-m})\varepsilon_t$ $s_t = s_{t-m} + \gamma(\ell_{t-1} + b_{t-1} + s_{t-m})\varepsilon_t$	$y_t = (\ell_{t-1} + b_{t-1})s_{t-m}(1 + \varepsilon_t)$ $\ell_t = (\ell_{t-1} + b_{t-1})(1 + \alpha\varepsilon_t)$ $b_t = b_{t-1} + \beta(\ell_{t-1} + b_{t-1})\varepsilon_t$ $s_t = s_{t-m}(1 + \gamma\varepsilon_t)$
A_d	$y_t = (\ell_{t-1} + \phi b_{t-1})(1 + \varepsilon_t)$ $\ell_t = (\ell_{t-1} + \phi b_{t-1})(1 + \alpha\varepsilon_t)$ $b_t = \phi b_{t-1} + \beta(\ell_{t-1} + \phi b_{t-1})\varepsilon_t$	$y_t = (\ell_{t-1} + \phi b_{t-1} + s_{t-m})(1 + \varepsilon_t)$ $\ell_t = \ell_{t-1} + \phi b_{t-1} + \alpha(\ell_{t-1} + \phi b_{t-1} + s_{t-m})\varepsilon_t$ $b_t = \phi b_{t-1} + \beta(\ell_{t-1} + \phi b_{t-1} + s_{t-m})\varepsilon_t$ $s_t = s_{t-m} + \gamma(\ell_{t-1} + \phi b_{t-1} + s_{t-m})\varepsilon_t$	$y_t = (\ell_{t-1} + \phi b_{t-1})s_{t-m}(1 + \varepsilon_t)$ $\ell_t = (\ell_{t-1} + \phi b_{t-1})(1 + \alpha\varepsilon_t)$ $b_t = \phi b_{t-1} + \beta(\ell_{t-1} + \phi b_{t-1})\varepsilon_t$ $s_t = s_{t-m}(1 + \gamma\varepsilon_t)$

- Smoothing parameters α , β , γ and ϕ , and the initial states l_0 , b_0 , s_0 , s_{-1} , \dots , s_{-m+1} are estimated by maximising the “likelihood” = the probability of the data arising from the specified model.
- For models with additive errors equivalent to minimising SSE.
- For models with multiplicative errors, **not** equivalent to minimising SSE.

Innovations state space models

Let $\mathbf{x}_t = (\ell_t, \mathbf{b}_t, s_t, s_{t-1}, \dots, s_{t-m+1})$ and $\varepsilon_t \stackrel{\text{iid}}{\sim} \mathcal{N}(0, \sigma^2)$.

$$y_t = \underbrace{h(\mathbf{x}_{t-1})}_{\mu_t} + \underbrace{k(\mathbf{x}_{t-1})\varepsilon_t}_{e_t}$$

$$\mathbf{x}_t = f(\mathbf{x}_{t-1}) + g(\mathbf{x}_{t-1})\varepsilon_t$$

Additive errors

$$k(\mathbf{x}) = 1. \quad y_t = \mu_t + \varepsilon_t.$$

Multiplicative errors

$$k(\mathbf{x}_{t-1}) = \mu_t. \quad y_t = \mu_t(1 + \varepsilon_t).$$

$\varepsilon_t = (y_t - \mu_t)/\mu_t$ is relative error.

Estimation

$$\begin{aligned}L^*(\boldsymbol{\theta}, \mathbf{x}_0) &= T \log \left(\sum_{t=1}^T \varepsilon_t^2 \right) + 2 \sum_{t=1}^T \log |k(\mathbf{x}_{t-1})| \\ &= -2 \log(\text{Likelihood}) + \text{constant}\end{aligned}$$

- Estimate parameters $\boldsymbol{\theta} = (\alpha, \beta, \gamma, \phi)$ and initial states $\mathbf{x}_0 = (\ell_0, b_0, s_0, s_{-1}, \dots, s_{-m+1})$ by minimizing L^* .

Usual region

- Traditional restrictions in the methods $0 < \alpha, \beta^*, \gamma^*, \phi < 1$ (equations interpreted as weighted averages).
- In models we set $\beta = \alpha\beta^*$ and $\gamma = (1 - \alpha)\gamma^*$.
- Therefore $0 < \alpha < 1$, $0 < \beta < \alpha$ and $0 < \gamma < 1 - \alpha$.
- $0.8 < \phi < 0.98$ — to prevent numerical difficulties.

Parameter restrictions

Usual region

- Traditional restrictions in the methods $0 < \alpha, \beta^*, \gamma^*, \phi < 1$ (equations interpreted as weighted averages).
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- Therefore $0 < \alpha < 1$, $0 < \beta < \alpha$ and $0 < \gamma < 1 - \alpha$.
- $0.8 < \phi < 0.98$ — to prevent numerical difficulties.

Admissible region

- To prevent observations in the distant past having a continuing effect on current forecasts.
- Usually (but not always) less restrictive than *traditional* region.
- For example for ETS(A,N,N):
traditional $0 < \alpha < 1$ while *admissible* $0 < \alpha < 2$.

Akaike's Information Criterion

$$\text{AIC} = -2 \log(L) + 2k$$

where L is the likelihood and k is the number of parameters initial states estimated in the model.

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Corrected AIC

$$\text{AIC}_c = \text{AIC} + \frac{2k(k+1)}{T-k-1}$$

which is the AIC corrected (for small sample bias).

Model selection

Akaike's Information Criterion

$$\text{AIC} = -2 \log(L) + 2k$$

where L is the likelihood and k is the number of parameters initial states estimated in the model.

Corrected AIC

$$\text{AIC}_c = \text{AIC} + \frac{2k(k+1)}{T-k-1}$$

which is the AIC corrected (for small sample bias).

Bayesian Information Criterion

$$\text{BIC} = \text{AIC} + k[\log(T) - 2].$$

Minimizing the AIC assuming Gaussian residuals is asymptotically equivalent to minimizing one-step time series cross validation MSE.

From Hyndman et al. (IJF, 2002):

- Apply each model that is appropriate to the data. Optimize parameters and initial values using MLE (or some other criterion).
- Select best method using AICc:
- Produce forecasts using best method.
- Obtain forecast intervals using underlying state space model.

Method performed very well in M3 competition.

Example: National populations

```
fit <- global_economy %>%  
  mutate(Pop = Population / 1e6) %>%  
  model(ets = ETS(Pop))  
fit
```

```
## # A mtable: 263 x 2  
## # Key:      Country [263]  
##   Country                ets  
##   <fct>                  <model>  
## 1 Afghanistan            <ETS(A,A,N)>  
## 2 Albania                 <ETS(M,A,N)>  
## 3 Algeria                 <ETS(M,A,N)>  
## 4 American Samoa         <ETS(M,A,N)>  
## 5 Andorra                 <ETS(M,A,N)>  
## 6 Angola                  <ETS(M,A,N)>  
## 7 Antigua and Barbuda    <ETS(M,A,N)>  
## 8 Arab World              <ETS(M,A,N)>
```

Example: National populations

```
fit %>%  
  forecast(h = 5)
```

```
## # A fable: 1,315 x 5 [1Y]  
## # Key:      Country, .model [263]  
##   Country      .model  Year          Pop .mean  
##   <fct>        <chr>  <dbl>         <dist> <dbl>  
## 1 Afghanistan ets     2018    N(36, 0.012) 36.4  
## 2 Afghanistan ets     2019    N(37, 0.059) 37.3  
## 3 Afghanistan ets     2020    N(38, 0.16)  38.2  
## 4 Afghanistan ets     2021    N(39, 0.35)  39.0  
## 5 Afghanistan ets     2022    N(40, 0.64)  39.9  
## 6 Albania      ets     2018    N(2.9, 0.00012) 2.87  
## 7 Albania      ets     2019    N(2.9, 6e-04)  2.87  
## 8 Albania      ets     2020    N(2.9, 0.0017) 2.87  
## 9 Albania      ets     2021    N(2.9, 0.0036) 2.86
```

Example: Australian holiday tourism

```
holidays <- tourism %>%  
  filter(Purpose == "Holiday")  
fit <- holidays %>% model(ets = ETS(Trips))  
fit
```

```
## # A mtable: 76 x 4
```

```
## # Key:      Region, State, Purpose [76]
```

##	Region	State	Purpose	ets
##	<chr>	<chr>	<chr>	<model>
##	1 Adelaide	South Australia	Holiday	<ETS(A,N,A)>
##	2 Adelaide Hills	South Australia	Holiday	<ETS(A,A,N)>
##	3 Alice Springs	Northern Territory	Holiday	<ETS(M,N,A)>
##	4 Australia's Coral Coast	Western Australia	Holiday	<ETS(M,N,A)>
##	5 Australia's Golden Outback	Western Australia	Holiday	<ETS(M,N,M)>
##	6 Australia's North West	Western Australia	Holiday	<ETS(A,N,A)>
##	7 Australia's South West	Western Australia	Holiday	<ETS(M,N,M)>
##	8 Ballarat	Victoria	Holiday	<ETS(M,N,A)>

Example: Australian holiday tourism

```
fit %>%  
  filter(Region == "Snowy Mountains") %>%  
  report()
```

```
## Series: Trips  
## Model: ETS(M,N,A)  
## Smoothing parameters:  
##   alpha = 0.157  
##   gamma = 1e-04  
##  
## Initial states:  
## l[0] s[0] s[-1] s[-2] s[-3]  
## 142 -61 131 -42.2 -27.7  
##  
## sigma^2: 0.0388  
##  
## AIC AICc BIC
```

Example: Australian holiday tourism

```
fit %>%  
  filter(Region == "Snowy Mountains") %>%  
  components(fit)
```

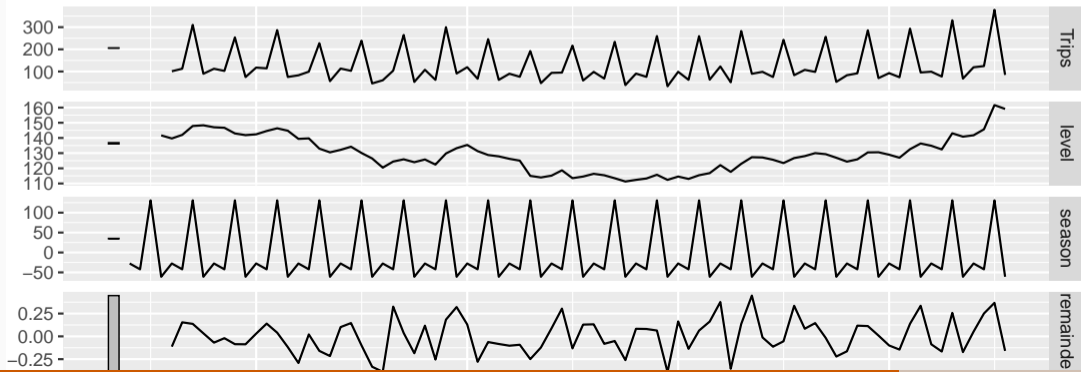
```
## # A dtable: 84 x 9 [1Q]  
## # Key:      Region, State, Purpose, .model [1]  
## # :      Trips = (lag(level, 1) + lag(season, 4)) * (1 + remainder)  
##   Region      State Purpose .model Quarter Trips level season remainder  
##   <chr>        <chr> <chr>  <chr>   <qtr> <dbl> <dbl> <dbl>    <dbl>  
## 1 Snowy Mountai~ New ~ Holiday ets    1997 Q1  NA     NA    -27.7   NA  
## 2 Snowy Mountai~ New ~ Holiday ets    1997 Q2  NA     NA    -42.2   NA  
## 3 Snowy Mountai~ New ~ Holiday ets    1997 Q3  NA     NA    131.    NA  
## 4 Snowy Mountai~ New ~ Holiday ets    1997 Q4  NA    142.   -61.0   NA  
## 5 Snowy Mountai~ New ~ Holiday ets    1998 Q1 101.   140.   -27.7  -0.113  
## 6 Snowy Mountai~ New ~ Holiday ets    1998 Q2 112.   142.   -42.2   0.154  
## 7 Snowy Mountai~ New ~ Holiday ets    1998 Q3 310.   148.   131.    0.137  
## 8 Snowy Mountai~ New ~ Holiday ets    1998 Q4  99.9   148.   -61.0   0.2225
```

Example: Australian holiday tourism

```
fit %>%  
  filter(Region == "Snowy Mountains") %>%  
  components(fit) %>%  
  autoplot()
```

ETS(M,N,A) decomposition

Trips = (lag(level, 1) + lag(season, 4)) * (1 + remainder)



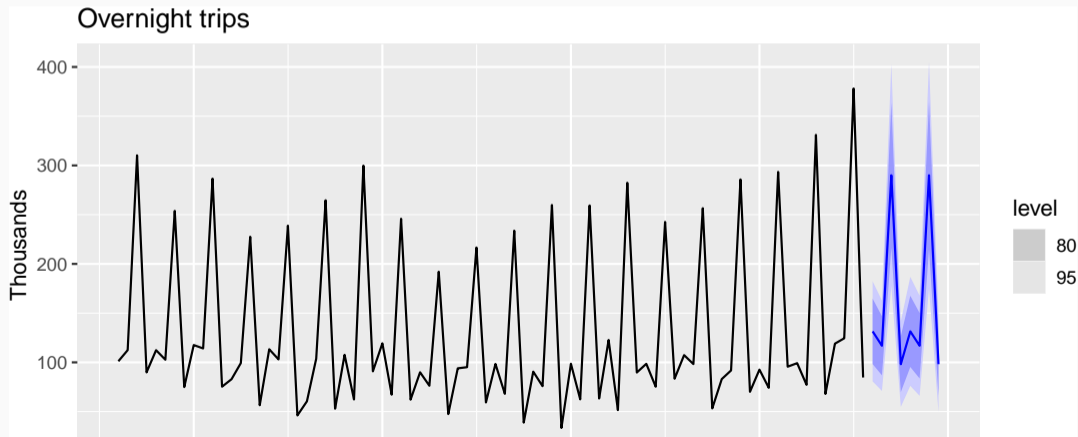
Example: Australian holiday tourism

```
fit %>% forecast()
```

```
## # A fable: 608 x 7 [1Q]
## # Key:      Region, State, Purpose, .model [76]
##   Region      State      Purpose .model Quarter      Trips .mean
##   <chr>        <chr>        <chr>  <chr>    <qtr>      <dist> <dbl>
## 1 Adelaide     South Australia Holiday ets    2018 Q1 N(210, 457) 210.
## 2 Adelaide     South Australia Holiday ets    2018 Q2 N(173, 473) 173.
## 3 Adelaide     South Australia Holiday ets    2018 Q3 N(169, 489) 169.
## 4 Adelaide     South Australia Holiday ets    2018 Q4 N(186, 505) 186.
## 5 Adelaide     South Australia Holiday ets    2019 Q1 N(210, 521) 210.
## 6 Adelaide     South Australia Holiday ets    2019 Q2 N(173, 537) 173.
## 7 Adelaide     South Australia Holiday ets    2019 Q3 N(169, 553) 169.
## 8 Adelaide     South Australia Holiday ets    2019 Q4 N(186, 569) 186.
## 9 Adelaide Hills South Australia Holiday ets    2018 Q1  N(19, 36)  19.4
## 10 Adelaide Hills South Australia Holiday ets    2018 Q2  N(20, 36)  19.6
```

Example: Australian holiday tourism

```
fit %>% forecast() %>%  
  filter(Region == "Snowy Mountains") %>%  
  autoplot(holidays) +  
  labs(y = "Thousands", title = "Overnight trips")
```



Response residuals

$$\hat{e}_t = y_t - \hat{y}_{t|t-1}$$

Innovation residuals

Additive error model:

$$\hat{\varepsilon}_t = y_t - \hat{y}_{t|t-1}$$

Multiplicative error model:

$$\hat{\varepsilon}_t = \frac{y_t - \hat{y}_{t|t-1}}{\hat{y}_{t|t-1}}$$

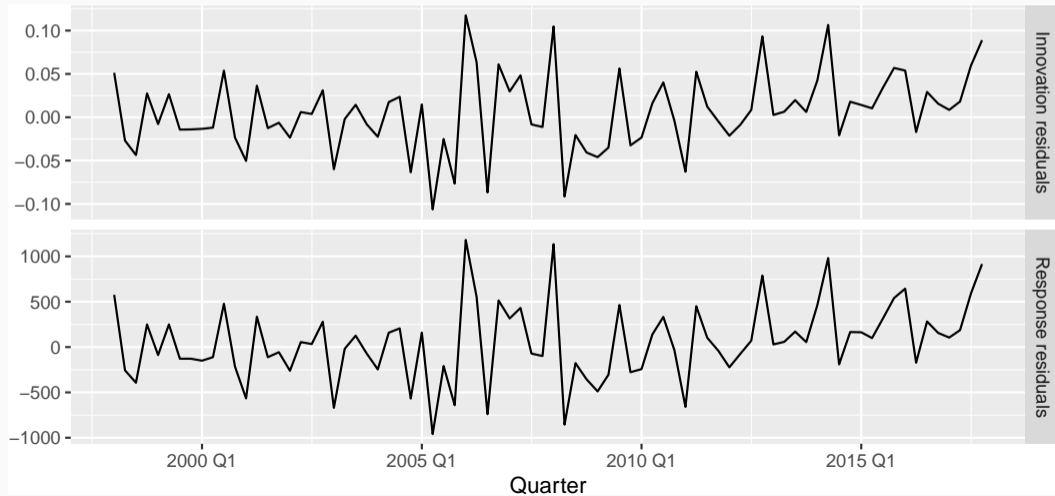
Example: Australian holiday tourism

```
aus_holidays <- tourism %>%  
  filter(Purpose == "Holiday") %>%  
  summarise(Trips = sum(Trips))  
fit <- aus_holidays %>%  
  model(ets = ETS(Trips)) %>%  
  report()
```

```
## Series: Trips  
## Model: ETS(M,N,M)  
## Smoothing parameters:  
##   alpha = 0.358  
##   gamma = 0.000969  
##  
## Initial states:  
## l[0] s[0] s[-1] s[-2] s[-3]  
## 9667 0.943 0.927 0.968 1.16  
##
```

Example: Australian holiday tourism

```
residuals(fit)  
residuals(fit, type = "response")
```



Example: Australian holiday tourism

```
fit %>%  
  augment()
```

```
## # A tsibble: 80 x 6 [1Q]  
## # Key:       .model [1]  
##   .model Quarter  Trips .fitted .resid  .innov  
##   <chr>    <qtr> <dbl> <dbl> <dbl> <dbl>  
## 1 ets     1998 Q1 11806. 11230.  576.  0.0513  
## 2 ets     1998 Q2  9276.  9532. -257. -0.0269  
## 3 ets     1998 Q3  8642.  9036. -393. -0.0435  
## 4 ets     1998 Q4  9300.  9050.  249.  0.0275  
## 5 ets     1999 Q1 11172. 11260. -88.0 -0.00781  
## 6 ets     1999 Q2  9608.  9358.  249.  0.0266  
## 7 ets     1999 Q3  8914.  9042. -129. -0.0142  
## 8 ets     1999 Q4  9026.  9154. -129. -0.0140  
## 9 ets     2000 Q1 11071. 11221. -150. -0.0134
```

Example: Australian holiday tourism

```
fit %>%  
  augment()
```

```
## # A tibble: 80 x 6 [1Q]  
## # Key:       .model [1]  
##   .model Quarter  Trips .fitted     .resid     .innov  
##   <chr>    <qtr> <dbl> <dbl> <dbl> <dbl>  
## 1 ets     1998 Q1 11806. 11230.  576.  0.0513  
## 2 ets     1998 Q2  9276.  9532. -257. -0.0269  
## 3 ets     1998 Q3  8642.  9036. -393. -0.0435  
## 4 ets     1998 Q4  9300.  9050.  249.  0.0275  
## 5 ets     1999 Q1 11172. 11260. -88.0 -0.00781  
## 6 ets     1999 Q2  9608.  9358.  249.  0.0266  
## 7 ets     1999 Q3  8914.  9042. -129. -0.0142  
## 8 ets     1999 Q4  9026.  9154. -129. -0.0140  
## 9 ets     2000 Q1 11071. 11221. -150. -0.0134
```

Innovation residuals (`.innov`) are given by $\hat{\varepsilon}_t$ while regular residuals (`.resid`) are $y_t - \hat{y}_{t-1}$. They are different when the model has multiplicative errors.

- Some of the combinations of (Error, Trend, Seasonal) can lead to numerical difficulties; see equations with division by a state.
- These are: $ETS(A,N,M)$, $ETS(A,A,M)$, $ETS(A,A_d,M)$.
- Models with multiplicative errors are useful for strictly positive data, but are not numerically stable with data containing zeros or negative values. In that case only the six fully additive models will be applied.

Exponential smoothing models

Additive Error

Trend Component		Seasonal Component		
		N (None)	A (Additive)	M (Multiplicative)
N	(None)	A,N,N	A,N,A	A,N,M
A	(Additive)	A,A,N	A,A,A	A,A,M
A _d	(Additive damped)	A,A _d ,N	A,A _d ,A	A,A_d,M

Multiplicative Error

Trend Component		Seasonal Component		
		N (None)	A (Additive)	M (Multiplicative)
N	(None)	M,N,N	M,N,A	M,N,M
A	(Additive)	M,A,N	M,A,A	M,A,M
A _d	(Additive damped)	M,A _d ,N	M,A _d ,A	M,A _d ,M

- 1 Exponential smoothing
- 2 Simple exponential smoothing
- 3 Models with trend
- 4 Models with seasonality
- 5 Innovations state space models
- 6 Forecasting with exponential smoothing

Traditional point forecasts: iterate the equations for $t = T + 1, T + 2, \dots, T + h$ and set all $\varepsilon_t = 0$ for $t > T$.

Traditional point forecasts: iterate the equations for $t = T + 1, T + 2, \dots, T + h$ and set all $\varepsilon_t = 0$ for $t > T$.

- Not the same as $E(y_{t+h} | \mathbf{x}_t)$ unless seasonality is additive.
- `fable` uses $E(y_{t+h} | \mathbf{x}_t)$.
- Point forecasts for $\text{ETS}(A, *, *)$ are identical to $\text{ETS}(M, *, *)$ if the parameters are the same.

$$y_{T+1} = l_T + b_T + \varepsilon_{T+1}$$

$$\hat{y}_{T+1|T} = l_T + b_T$$

$$\begin{aligned} y_{T+2} &= l_{T+1} + b_{T+1} + \varepsilon_{T+2} \\ &= (l_T + b_T + \alpha\varepsilon_{T+1}) + (b_T + \beta\varepsilon_{T+1}) + \varepsilon_{T+2} \end{aligned}$$

$$\hat{y}_{T+2|T} = l_T + 2b_T$$

etc.

$$y_{T+1} = (l_T + b_T)(1 + \varepsilon_{T+1})$$

$$\hat{y}_{T+1|T} = l_T + b_T.$$

$$y_{T+2} = (l_{T+1} + b_{T+1})(1 + \varepsilon_{T+2})$$

$$= \{(l_T + b_T)(1 + \alpha\varepsilon_{T+1}) + [b_T + \beta(l_T + b_T)\varepsilon_{T+1}]\} (1 + \varepsilon_{T+2})$$

$$\hat{y}_{T+2|T} = l_T + 2b_T$$

etc.

Prediction intervals: can only be generated using the models.

- The prediction intervals will differ between models with additive and multiplicative errors.
- Exact formulae for some models.
- More general to simulate future sample paths, conditional on the last estimate of the states, and to obtain prediction intervals from the percentiles of these simulated future paths.

Prediction intervals

PI for most ETS models: $\hat{y}_{T+h|T} \pm c\sigma_h$, where c depends on coverage probability and σ_h is forecast standard deviation.

$$(A,N,N) \quad \sigma_h = \sigma^2 \left[1 + \alpha^2(h-1) \right]$$

$$(A,A,N) \quad \sigma_h = \sigma^2 \left[1 + (h-1) \left\{ \alpha^2 + \alpha\beta h + \frac{1}{6}\beta^2 h(2h-1) \right\} \right]$$

$$(A,A_d,N) \quad \sigma_h = \sigma^2 \left[1 + \alpha^2(h-1) + \frac{\beta\phi h}{(1-\phi)^2} \{2\alpha(1-\phi) + \beta\phi\} - \frac{\beta\phi(1-\phi^h)}{(1-\phi)^2(1-\phi^2)} \{2\alpha(1-\phi^2) + \beta\phi(1+2\phi-\phi^h)\} \right]$$

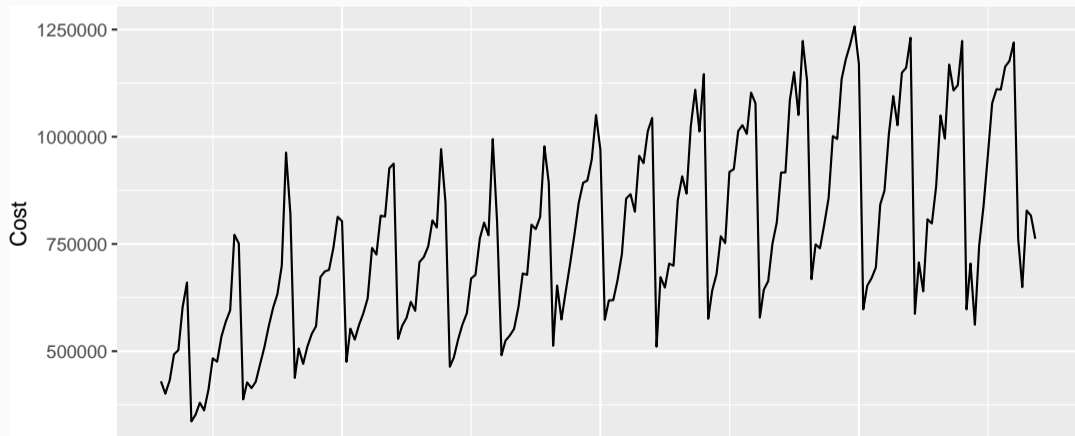
$$(A,N,A) \quad \sigma_h = \sigma^2 \left[1 + \alpha^2(h-1) + \gamma k(2\alpha + \gamma) \right]$$

$$(A,A,A) \quad \sigma_h = \sigma^2 \left[1 + (h-1) \left\{ \alpha^2 + \alpha\beta h + \frac{1}{6}\beta^2 h(2h-1) \right\} + \gamma k \{2\alpha + \gamma + \beta m(k+1)\} \right]$$

$$(A,A_d,A) \quad \sigma_h = \sigma^2 \left[1 + \alpha^2(h-1) + \frac{\beta\phi h}{(1-\phi)^2} \{2\alpha(1-\phi) + \beta\phi\} - \frac{\beta\phi(1-\phi^h)}{(1-\phi)^2(1-\phi^2)} \{2\alpha(1-\phi^2) + \beta\phi(1+2\phi-\phi^h)\} \right. \\ \left. + \gamma k(2\alpha + \gamma) + \frac{2\beta\gamma\phi}{(1-\phi)(1-\phi^m)} \{k(1-\phi^m) - \phi^m(1-\phi^{mk})\} \right]$$

Example: Corticosteroid drug sales

```
h02 <- PBS %>%  
  filter(ATC2 == "H02") %>%  
  summarise(Cost = sum(Cost))  
h02 %>% autoplot(Cost)
```



Example: Corticosteroid drug sales

```
h02 %>%
```

```
  model(ETS(Cost)) %>%
```

```
  report()
```

```
## Series: Cost
```

```
## Model: ETS(M,Ad,M)
```

```
## Smoothing parameters:
```

```
##   alpha = 0.307
```

```
##   beta  = 0.000101
```

```
##   gamma = 0.000101
```

```
##   phi   = 0.978
```

```
##
```

```
## Initial states:
```

```
##   l[0] b[0] s[0] s[-1] s[-2] s[-3] s[-4] s[-5] s[-6] s[-7] s[-8] s[-9]
```

```
## 417269 8206 0.872 0.826 0.756 0.773 0.687 1.28 1.32 1.18 1.16 1.1
```

```
## s[-10] s[-11]
```

```
## 1.05 0.001
```

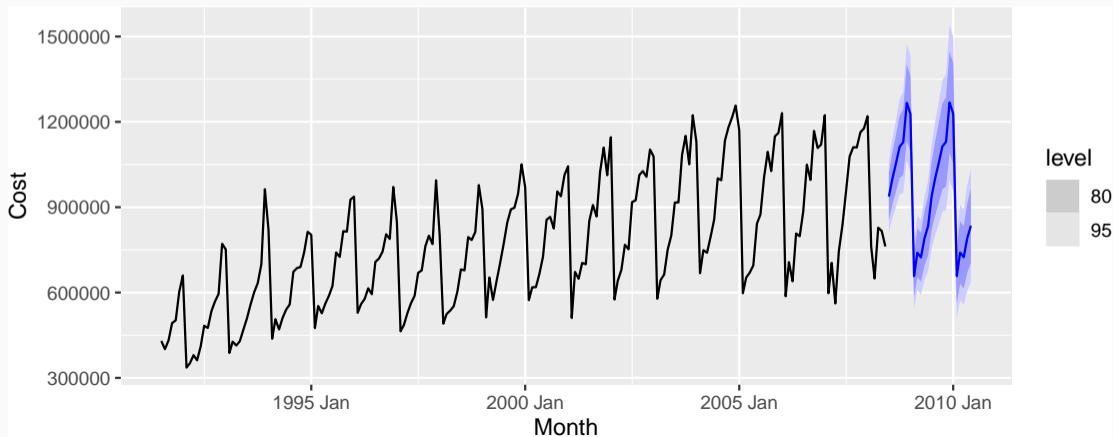
Example: Corticosteroid drug sales

```
h02 %>%  
  model(ETS(Cost ~ error("A") + trend("A") + season("A"))) %>%  
  report()
```

```
## Series: Cost  
## Model: ETS(A,A,A)  
## Smoothing parameters:  
##   alpha = 0.17  
##   beta  = 0.00631  
##   gamma = 0.455  
##  
## Initial states:  
##   l[0] b[0]   s[0]   s[-1]   s[-2]   s[-3]   s[-4]   s[-5]   s[-6]   s[-7]  
## 409706 9097 -99075 -136602 -191496 -174531 -241437 210644 244644 145368  
##   s[-8] s[-9] s[-10] s[-11]  
## 130570 84458 39132 -11674  
##
```

Example: Corticosteroid drug sales

```
h02 %>%  
  model(ETS(Cost)) %>%  
  forecast() %>%  
  autoplot(h02)
```



Example: Corticosteroid drug sales

```
h02 %>%  
  model(  
    auto = ETS(Cost),  
    AAA = ETS(Cost ~ error("A") + trend("A") + season("A"))  
  ) %>%  
  accuracy()
```

Model	MAE	RMSE	MAPE	MASE	RMSSE
auto	38649	51102	4.99	0.638	0.689
AAA	43378	56784	6.05	0.716	0.766